

Long-Term Asset Price Volatility and Macroeconomic Fluctuations

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We analyze a stochastic growth model with lags in the operation of new technologies. Stock values are impacted by news on technological innovations and some other external forces affecting the economy. Episodes of technology adoption may generate long fluctuations in the aggregate value of stocks. We assess the quantitative importance of various macroeconomic variables such as markups, leverage, adjustment costs, taxes, and labor and financial frictions. The model can account for a sizable part of the volatility observed in stock values while preserving the volatilities of real macroeconomic aggregates.

KEY WORDS: Technological innovation, stock market, markups, leverage, taxes, labor and financial frictions.

1 Introduction

This paper is concerned with macroeconomic determinants of the volatility of asset prices. Stochastic growth models have been fairly successful in accounting for co-movements of real

economic aggregates but have failed to offer plausible explanations for the volatility of stock values based on the variability of economic fundamentals. Hence, the modern literature in the interface between macroeconomics and finance has consistently appealed to several forms of sophisticated preferences, stochastic discounting, time varying risk, adjustment costs, bubbles, and noise trader risk. We shall here take a more traditional route, and work with a standard intertemporal preference representation under a CES utility function and constant discounting, and a mild form of adjustment costs for capital investment. Our main purpose here is to build a quantitative framework that can account for a good part of the observed volatility in stock market values along with the much lower volatility of real economic aggregates.

In the neoclassical growth model, changes in total factor productivity (TFP), taxes, the relative price of capital, or frictions in labor and capital markets hardly generate any volatility of stock values [cf. Rouwenhorst (1995)]. What happens in this model is that these forces do not affect significantly the volatility and persistence of dividends and earnings under observable variations in consumption. Viewed in another way, capital is the only asset in the economy, and investment must fluctuate enormously to get desirable levels of volatility in stock values. The model's performance can be improved with adjustment costs [Christiano and Fisher (2003)], but these costs may seem implausibly high.

We consider a stochastic growth model with lags in the operation of new technologies. Technological innovations arrive exogenously to the economy and impact stocks through their option values. These innovations, however, cannot be readily put to use and undergo a process of adoption embedded in the production of new varieties of intermediate goods. Episodes of technology adoption may generate long fluctuations in the aggregate value of stocks. An important consideration is that asset prices incorporate the option value of technological innovations that have yet to be implemented. This propagation mechanism is likewise present in the partial equilibrium setting of Abel and Eberly (2005), and in the tree economy of Panageas and Yu (2006). Hence, the value of a firm may differ from the replacement value of its durable factors of production.

There is a burgeoning literature on macroeconomics and finance, but we still lack a good understanding of the main macroeconomic determinants of asset price volatility. A good part of the literature has focused on the equity premium [e.g., see Campbell (1999), Cochrane and

Hansen (1992) and Mehra and Prescott (2003)]. There is much less work intended to offer a joint explanation for the volatilities of the stock market and other macroeconomic variables [cf., Backus, Routledge and Zin (2006), Christiano and Fisher (2003), and Rouwenhorst (1995)]. In a recent paper, Jinnai (2009) argues in favor of two-sector models and evaluates the influence of intangible capital, leverage, habit formation, and adjustment costs on asset volatility and the equity premium.

Our model is a simplified variant of those in Romer (1990) and Comin and Gertler (2006), but our objectives are quite different. Romer (1990) is concerned with innovations and economic growth and Comin and Gertler (2006) with a quantitative analysis of economic fluctuations.¹ Our main goal is to examine possible sources of volatility of stock market prices while preserving the less pronounced volatility of real macroeconomic aggregates.

At a later stage we report some quantitative experiments. In spite of restricting the analysis to a simple CES utility function and a constant discount factor, our model can account for a sizable part of the volatility of stock market values. The model is solved numerically by a high-order approximation method that picks non-linearities in the evolution of stock prices. We assign parameter values to fit some basic facts in economic growth and business fluctuations. Then, we perform several numerical exercises to see how innovations to the economy may affect the dynamics of stock prices and other aggregate variables. As in Greenwood and Jovanovic (1999) not all technological innovations will increase stock market prices, since the arrival of new technologies will depreciate the value of existing ones. To affect positively the stock market, new technologies must command higher price markups. Apart from price markups, we also consider the effects of leverage, adjustment costs, taxes, and labor and financial frictions. A notable feature of these exercises is that adjustment costs to capital investment play a minor role, and taxes and labor and financial frictions do not seem to affect substantially the volatility of asset values.

The paper will proceed as follows. We start in Section 2 with some empirical evidence on stock market fluctuations. In Section 3 we lay out our model of technology adoption and

¹In a later paper, Comin, Gertler and Santacreu (2009) borrow extensively from our theoretical work and build their main intuition and economic ideas from our fundamental equation for asset pricing derived in Proposition 3.1 below. It seems that most of the shocks that they consider have a small influence on the volatility of asset markets. Their results could be driven by stochastic discounting and a strong form of adjustment costs, rather than by the main forces underlying our quantitative experiments that yield much more volatility of dividends.

derive some qualitative properties of the solution with emphasis on a fundamental asset pricing equation that decomposes the stock price into the stock of physical capital and the value of adopted and unadopted technologies for the production of intermediate products. Section 4 is devoted to the computation and calibration of the model, and Section 5 reports various numerical experiments. We conclude in Section 6 with a further evaluation of our main findings.

2 Some Empirical Evidence

Figure 1 plots the evolution of the S&P and the corresponding price-earnings (PE) ratio for the period 1881–2005. These data have been filtered by taking out our best fit for a deterministic exponential trend. Observe that both series display similar long-term cyclical behavior; indeed, as discussed below stock prices are a main driving force of these fluctuations. Since the early papers of Kleidon (1986), Marsh and Merton (1987) and others, researchers have emphasized the persistence and lower variability of dividends and earnings, which are highly correlated. After some adjusting for interest and growth rates, Barsky and De Long (1993) obtain that extrapolations of dividends over twenty-year periods can mimic reasonably well the volatility of stocks. Figure 2 portrays a centered ten-year moving average of S&P yearly returns. We observe that peak values occur in 1880, 1900, 1925, 1955 and 1995. Therefore, the amplitude of these long-term cycles can be up to 40 years.

Jovanovic and Rousseau (2001) associate these long fluctuations in the stock market with three technological revolutions: Electricity, World War II, and IT. These authors document long lags in the operation and diffusion of new technologies. Nicholas (2008) claims that innovation was a main driver of the stock market run-up of the late 1920s. To see the financial impact of the recent IT revolution, Figure 3 decomposes the market capitalization relative to GNP into the values of four different groups of companies: (i) The incumbents, (ii) Companies originating in 1982-1989, (iii) Companies originating in 1990-1994, and (iv) Companies originating after 1994. As one can see, most added stock value belongs to new corporations. In our model these newcomers would be reflected in the value added of local technology adopters. In later work, Jovanovic and Rousseau (2009) address the further issue of when new technologies are implemented by either incumbents or new companies.

However, our story is not only a story of technology adoption. As a matter of fact, there seem to be other trends associated with these cycles coming from factors such as markups, population, housing prices, and taxes. Figure 4 plots the price of farmland in England and the price of oil which move in tandem – property values leading changes in oil prices. Figure 5 illustrates that the S&P 500 is also a good leading indicator of oil prices. Given the cost structure of the oil industry, oil prices must be associated with the volatility and persistency of price markups which in turn may be generated by global forces in demand and supply [Dvir and Rogoff (2009)]. Markups are substantial in the US economy [Hall (1988a)]. Markups can vary over time led by macro trends and are prominent in the innovation and product cycles [Rotemberg and Woodford (1995) and Broda and Weinstein (2007)].

Geanakoplos, Magill and Quinzii (2004) contend that changes in stock values may be driven by demographic trends, whilst Lustig and Nieuwerburgh (2006) cite credit access from home equity collateral that may affect attitudes toward risk. In Geanakoplos, Magill and Quinzii (2004) the value of the stock market is determined by the ratio of middle aged to young people (MY). Because of the different saving habits of these two groups the stock market would be high when MY is high. Of course, this mechanism ignores the globalization of financial markets. Moreover, the introduction of uncertainty may have a dampening effect in these cycles, since the young generation is less risk averse and may want to hold a large part of the stock. Our model does not consider housing, and hence we cannot formally explore the propagation mechanisms in Lustig and Nieuwerburgh (2006). Nevertheless, it is not clear to us the direction of causality, since interest rates, stock and housing values are all endogenously determined. Moreover, Figure 6 highlights that high equity premia usually occur in periods of high returns for the stock market. These periods are generally economic expansions with high, possibly unexpected capital gains – where consumers are not especially cash constrained or risk averse. Therefore, it seems difficult to match these data with theories of habit formation, risk aversion, or access to credit. As a matter of fact, high excess returns are not highly correlated with consumption growth [Imrohoroglu (2003)].

Several recent works have focused on the collapse of stocks in the 70s and subsequent periods of recovery. A large body of research [Greenwood and Jovanovic (1999), Laitner and Stolyarov (2003) and Peralta-Alva (2006)] elaborates on the effects of IT on the dynamics

of capital values.² McGrattan and Prescott (2005) attribute these recent trends in stock values to changes in the tax system. All these papers are mainly concerned with levels effects, whereas we analyze the volatilities of both financial and real variables in a business cycle framework. In the same way, we interpret that most papers in monetary policy and collateral requirements [Bernanke and Gertler (1999), Christiano et al. (2008), and Kiyotaki and Moore (1997)] stress some important qualitative level effects of financial interventions but it remains to be explored that full-blown versions of these models can generate plausible volatilities for asset values.

3 The Model

The economy is populated by a continuum of identical households. At every time $t = 0, 1, \dots$, each agent demands quantities of the aggregate consumption good, supplies labor inelastically, and trades in the equity and bond markets. The aggregate consumption good is produced by a single firm with a constant returns to scale technology. Three inputs are involved in the production of this final commodity: Capital accumulated by the firm, labor, and a composite intermediate good. Both the firm and the consumer act competitively in all markets, but the sector of intermediate goods is composed of a continuum of monopolistic competitors. The range of intermediate goods produced can be expanded by a fixed set of local adopters upon the arrival of new technologies. An increase in the varieties of intermediate goods allows for a more efficient use of resources, and augments TFP. The remaining source of change in TFP is an exogenous shock to the production function of the final good. Our primary interest is to determine the most relevant forces affecting asset prices in this economy. Proposition 3.1 below puts forward an asset pricing equation which will be the basis for our later analysis.

3.1 *The household*

The representative household supplies labor inelastically, and has preferences over infinite streams of consumption. Preferences are represented by the expected discounted objective:

²See Boldrin and Levine (2001) and Hobbijn and Jovanovic (2001) for related models of technology adoption.

$$E_0 \left\{ \sum_{t=0}^{\infty} \beta^t \frac{c_t^{1-\sigma}}{1-\sigma} \right\} \quad (1)$$

where c_t denotes the quantity of consumption at t , with $0 < \beta < 1$ and $\sigma \geq 0$. To purchase a consumption plan, an agent may participate in financial markets by trading shares of an aggregate stock a_t and units of a risk-free bond b_t . The aggregate stock yields a stochastic aggregate dividend d_t comprising the pay-outs from the aforementioned production sectors, and the bond sells at a predetermined interest rate r_t . For given initial positive asset holdings \hat{a}, \hat{b} , the optimization problem of a representative agent is to choose a stochastic sequence of consumption, shares of the aggregate stock, and units of the risk-free bond $\{c_t, a_t, b_t\}_{t \geq 0}$ to attain the maximum utility in (1) subject to the sequence of budget constraints

$$c_t + q_t a_t + b_t = \omega_t l_t + (q_t + d_t) a_{t-1} + r_t b_{t-1} \quad (2)$$

$$q_t a_t + b_t \geq 0, \quad t = 0, 1, 2, \dots \quad (3)$$

for a given sequence of stock prices q_t , rates of interest r_t , and unitary wages ω_t . Note that in this representative agent economy the nonnegativity condition on portfolio holdings (3) entails no loss of generality, but of course this restriction can be relaxed.

3.2 The production sector

The firm producing the final good accumulates capital and buys labor and intermediate goods. TFP of the firm is stochastic, and represented by a random variable θ_t . At every date t there is a mass of A_t intermediate goods that enter into the production of the final good. These intermediate goods are bundled together in a composite good M_t defined by a CES technology, $M_t = [\int_0^{A_t} m_{s,t}^{\frac{1}{\vartheta_t}} ds]^{\vartheta_t}$ where $m_{s,t}$ denotes the amount of intermediate good s bought by the firm at time t and $\vartheta_t > 1$ follows an exogenous stochastic process to be specified below.

Given initial levels of capital and debt $\hat{k}, \hat{B} > 0$, the firm chooses stochastic sequences of investment, labor, and intermediate goods $\{i_t, l_t, (m_{s,t})_{s \in [0, A_t]}\}_{t \geq 0}$ so as to maximize the present value of dividends:

$$E_0 \left\{ \sum_{t=0}^{\infty} \eta_t d_t^f \right\} \quad (4)$$

subject to

$$d_t^f \equiv Y_t - i_t - \omega_t l_t - \int_0^{A_t} p_{s,t} m_{s,t} ds + B_t - r_t B_{t-1} \quad (5)$$

$$Y_t \equiv \theta_t \left[\gamma (k_t^\alpha l_t^{1-\alpha})^\rho + (1-\gamma) M_t^\rho \right]^{\frac{1}{\rho}}, \quad 0 < \gamma < 1, \quad -\infty < \rho < 1 \quad (6)$$

$$k_{t+1} = (1-\delta) k_t + g(i_t/k_t) k_t, \quad \text{and} \quad B_t \leq \bar{B}_t. \quad (7)$$

Observe that η_t is a state price converting income of period t to period 0, and $p_{s,t}$ denotes the price of intermediate good s at time t . The physical capital stock depreciates at a constant rate $\delta \geq 0$. Capital accumulation is also subject to adjustment costs which are represented by function g . This latter function is positive and convex with $g(\delta) = \delta$ and $g'(\delta) = 1$. The curvature of this function limits the volatility of capital investment.

Note that our definition of dividends in (5) includes financial leverage. The amount of debt B_t is bounded by an exogenous borrowing constraint \bar{B}_t . Later on, we will consider exogenous borrowing constraints for both long-term and short-term debt. Hall (2001) points out that debt policies have been quite volatile: Pay-outs to debt holders have been fairly erratic in recent decades.

In addition to stock prices, interest rates and wages, the firm considers that TFP and price markups evolve exogenously. Stochastic variables θ_t and ϑ_t are governed by the following stationary first-order autoregressive process

$$\ln(\theta_t) = \psi^\theta \ln(\theta_{t-1}) + \varepsilon_t^\theta \quad (8)$$

$$\ln(\vartheta_t) = \psi_0^\vartheta + \psi_1^\vartheta \ln(\theta_{t-1}) + \varepsilon_t^\vartheta \quad (9)$$

where $\varepsilon_t^\theta, \varepsilon_t^\vartheta$ are random variables with $\varepsilon_t^\vartheta \geq 0$.

Monopolistic competition prevails in the market for intermediate goods. Each variety s belongs to a separate, independent producer. Without loss of generality, the production

process of each variety is given by the following simple form: One unit of good s requires only one unit of the final good. Then, producer of variety s picks an optimal pricing strategy $p_{s,t}$ and quantity $m_{s,t}$ from inspection of the downward-sloping demand for the product by the firm producing the aggregate commodity – after assuming a fixed set of prices and quantities for all other varieties. More precisely, for each time period t producer of variety s maximizes the amount of profits:

$$\pi_{s,t} \equiv \max_{m_{s,t}} \{p_{s,t}m_{s,t} - m_{s,t}\} \quad (10)$$

where $p_{s,t}$ depends on $m_{s,t}$, and can be read off from the the inverse demand function

$$p_{s,t} = \left(\frac{m_{s,t}}{M_t} \right)^{\frac{1-\vartheta_t}{\vartheta_t}} p_t \quad (11)$$

$$\text{and } p_t = \left(\int_0^{A_t} p_{s,t}^{\frac{1}{1-\vartheta_t}} ds \right)^{1-\vartheta_t}.$$

Production of intermediate goods may be discontinued because of exogenous factors. Let ϕ be the probability of survival of a technology at every date t . Then, the present value $V_{s,t}$ of operating technology s from the beginning of time t is defined as:

$$V_{s,t} = E_t \left\{ \sum_{r=t}^{\infty} \frac{\eta_r}{\eta_t} \phi^{r-t} \pi_{s,r} \right\}. \quad (12)$$

For simplicity, we assume that $\pi_{s,t}$ and $V_{s,t}$ are the same for all s .

3.3 Technology adoption

Technological innovations arrive exogenously to the economy. The total stock of technological innovations Z_t evolves according to the law of motion

$$Z_t = \phi Z_{t-1} + \mu x_t \quad (13)$$

with $\mu > 0$ and

$$\ln x_t = \psi^x \ln x_{t-1} + \varepsilon_t^x \quad \psi_x \in (0, 1), \quad \varepsilon_t^x \sim N(0, \sigma_x) \quad (14)$$

Technologies are put into use by local adopters. The adoption sector is composed of a continuum of agents $i \in [0, 1]$ that behave competitively. Each adopted technology sells at price V_t to a producer of intermediate goods. Let A_t^i the stock of already adopted technologies by agent i , and $\lambda(H_t^i)$ the probability of adopting a new technology after investing the amount of resources H_t^i . An adopter can deal with several projects, and hence we assume that the stock A_{t+1}^i follows the law of motion

$$A_{t+1}^i = \lambda(H_t^i)\phi [Z_t^i - A_t^i] + \phi A_t^i. \quad (15)$$

The optimal amount of expenditure H_t^i is derived from the following Bellman equation in which the value function is the option value J_t of a new technology:

$$J_t = \max_{H_t^i} \left\{ -H_t^i + \phi E_t \left[\frac{\eta_{t+1}}{\eta_t} (\lambda_t V_{t+1} + (1 - \lambda) J_{t+1}) \right] \right\}. \quad (16)$$

As is well known, this equation can be computed recursively by the method of successive approximations. It follows that the optimal amount of expenditure H_t^i is the same for all i . We then let the aggregate stock of adopted technologies $A_{t+1} = \int A_{t+1}^i di$.

3.4 Equilibrium and Asset Prices

In our model the exogenous state variables are the stock of available technologies Z_t , the addition of new varieties x_t , the TFP index θ_t , and the price markup ϑ_t . The endogenous state variables are the capital stock k_t , and the stock of adopted technologies A_t . The remaining variables are determined as solutions of the model from the above optimization problems, the market clearing and feasibility conditions, and the laws of motion of the exogenous state variables.

As suggested above, we adopt the convention that the stock market value includes all the above three production sectors. That is, $q_t a_t$ comprises the value of the objective (4) for the firm producing the final good, plus the discounted net value of profits over the set of intermediate goods and technology adoption. Hence, the aggregate dividend $d_t = d_t^f + \pi_t A_t - H_t(Z_t - A_t)$. In what follows we assume that the aggregate net supply of the asset equals one (i.e. $a_t = 1$) so that q_t corresponds to the value of the stock market. Therefore, market

clearing in the stock and bond markets requires $a_t = 1$, and $b_t = -B_t$. For the aggregate commodity, market clearing holds if

$$Y_t - A_t m_t = c_t + i_t + H_t(Z_t - A_t) \quad (17)$$

where Y_t denotes gross production of the final good and m_t is the quantity of each intermediate good produced. Hence, the production cost of the composite intermediate good is $A_t m_t$. Observe that the left-hand side of equation (17) is the value added generated in this economy, which is broken down into consumption, investment in physical capital, and investment in adopting new technologies.

The first order conditions for the representative household's correspond to the usual no-arbitrage conditions for the aggregate stock and the risk-free bond:

$$1 = E_t \left\{ \frac{\eta_{t+1}}{\eta_t} \left(\frac{d_{t+1} + q_{t+1}}{q_t} \right) \right\} \quad (18)$$

$$1 = E_t \left\{ \frac{\eta_{t+1}}{\eta_t} r_t \right\}. \quad (19)$$

Observe that the firm producing the final good will always demand positive amounts of each factor, and hence the first order-conditions under the objective in (4) will always hold with equality. Optimal positive expenditure in the adoption of new varieties requires:

$$1 = \lambda'_t \phi E_t \left\{ \frac{\eta_{t+1}}{\eta_t} (V_{t+1} - J_{t+1}) \right\} \quad (20)$$

Therefore, for a concave function $\lambda(H)$ optimal expenditure is positively correlated with the expected difference between the value of adopted and non-adopted varieties.

The following proposition is central to our study. It shows that the value of the stock market incorporates the value of adopted technologies and the option value to adopt new technologies (currently available or in the future).

Proposition 3.1 *In equilibrium the market value of the aggregate stock*

$$q_t = p_t^I k_{t+1} - B_t + V_t^+ A_t + J_t^+ (Z_t - A_t) + \xi_t \quad (21)$$

where $V_t^+ \equiv V_t - \pi_t$, $J_t^+ \equiv J_t + H_t$, $\xi_t \equiv E_t \left\{ \sum_{r=t+1}^{\infty} \frac{\eta_r}{\eta_t} J_r (Z_r - \phi Z_{r-1}) \right\}$, and $p_t^I \equiv \frac{1}{g_t}$.

Hence, the value of the stock market is the sum of five components: the replacement cost of installed capital, the amount of debt, the value of adopted technologies, the option value of inventions currently available but not yet implemented, and the present value of future inventions expected to happen. These latter components are further sources of volatility in the stock market over the stock of capital and the value of adopted technologies. We will analyze the dynamic evolution of these components under perturbations of the exogenous state variables as well as their correlation with real macro aggregates.

4 Calibration and Computation of the Model

The equilibrium is computed numerically through a high-order perturbation method [Schmitt-Grohé and Uribe (2004)] that takes into account the high volatility of stock market prices. To check for accuracy of the computed solution, we have combined this approximation method with a numerical dynamic programming algorithm [Santos (1999)] for the computation of Bellman's equation (16).

There are several ingredients in our calibration exercise, which is quite complex. First, various standard parameters are taken from the literature. These include parameters defining the utility function, the aggregate production function, and adjustment costs. Second, regarding parameter values of the sector of technology adoption [equation (16)] where there could be a wide range of empirical estimates, these parameters are selected to match some statistics of our model to the data. As a matter of fact, to avoid a very high sensitivity of optimal expenditure H , we postulate an expenditure function that becomes fairly parsimonious with the data. Third, for the estimation of the law of motion for TFP, markups, and technology innovation, we use a simulation-based estimation along the lines of Santos (2009). The outcome of this exercise is to obtain an optimal estimation of the covariance matrix for these three shocks, rather than correlating these shocks in an arbitrary manner. This is of independent interest as it shows how TFP, markups and technological innovation shocks are possibly correlated in the data. And fourth, within plausible empirical bounds of debt to equity ratios [e.g., see Hall (2001), Rouwenhorst (1995) and references therein] we perform various numerical experiments for arbitrary debt policies followed by the aggregate firm. Then, we settle down on a simple calibration that does not alter substantially the

volatility of dividends.

Some standard model's parameters are calibrated to match different statistics of medium-term and long run fluctuations observed in the data. Following Comin and Gertler (2006) we define medium-term cycles as those within a band of frequencies from 2 to 50 years. We then detrend the data using the filter of Christiano and Fitzgerald (2003). As our main objective is to analyze long-term fluctuations, we considered annual data. Output, and the Solow residual are for the non-farm business sector. The source is the Bureau of Labor Statistics (BLS). Consumption is measured as the sum of non-durables and services, and investment refers to non-residential structures. Both series are obtained from the Bureau of Economic Analysis (BEA). The stock price, one year interest rate, dividends, and earnings are from Robert Shiller's web page: <http://www.econ.yale.edu/shiller/data.htm>. Each variable is transformed in per-capita terms by dividing it by the population aged 15 to 64. The data span over the period between 1948 and 2004.

We assume no value to leisure. As is well known, standard RBC models do not generate enough volatility in worked hours [see Hansen and Wright (1992), and Kydland (1995)]. We could improve the performance of the model in this dimension by incorporating labor indivisibilities, or variable effort. However, these labor market refinements would not change significantly asset pricing volatility.

Following the RBC literature [see, e.g., Cooley and Prescott (1995)] we choose values for the set of parameters (β, α, δ) such that the deterministic steady state coincides with the average behavior of per-capita U.S. time series. Parameter β is fixed at 0.95, leading to an annual interest rate of 5.26%. It is an intermediate value between the mean real short term interest rate and the real rate of return on S&P price index: The mean short term interest rate over the period considered is 1.9% and the mean stock market rate of return is 9%. The parameter σ in the utility function is set to 4, which is between the range suggested by Hall (1988b). We set $\alpha = 0.33$ based on evidence of the average share of labor costs over total costs. The share of materials in gross output is assumed to be 0.5 in line with estimates for the manufacturing sector [see Jaimovich and Floetotto (2008)]. To match this ratio in the steady state we set $\gamma = 0.7$. Substituting the average investment to capital ratio into (7), implies a depreciation rate δ of 0.09 per year. The intermediate producers' gross markup in the deterministic steady state, is fixed at 1.1. This value is smaller than the lower bound

for the range of estimates of this parameter [Rotemberg and Woodford, (1995) provide an overview of microeconomic evidence]. However, the simulated mean for this parameter is 1.23 which belongs to the range of estimates.

We follow Jermann (1998) in the specification of the adjustment cost function:

$$g(i/k) = \frac{\delta^{\frac{1}{\varsigma}}}{1 - \frac{1}{\varsigma}} \left(\frac{i}{k}\right)^{1 - \frac{1}{\varsigma}} + \frac{\delta}{1 - \varsigma} \quad (22)$$

where the positive parameter ς is the elasticity of the investment to capital ratio with respect to Tobin's q . We set this parameter equal to 8 which is in line with Jinnai (2009). As we discuss below, our results present low sensitivity to this parameter.

The probability of adoption is determined by

$$\lambda(H_t) = \Lambda H_t^\kappa \quad (23)$$

with $\Lambda > 0$ and $\kappa \in (0, 1)$. We assume that this probability in the steady state is 0.2, and therefore the average adoption time is five years. This lag seems suitable see Mansfield (1989). The parameter κ is set to 0.95, implying that the steady state ratio of adoption expenditures over value added belongs to the interval defined by the ratio of R&D over GDP and development over GDP. The parameter ϕ determining the survival rate of each intermediate product, is set to 0.972 following empirical estimates [see Hall (2007)]

We assume that the ratio of adoption expenditures over output ($RH_t \equiv H_t(Z_t - A_t)/(Y_t - A_t m_t)$) satisfies the following expression

$$RH_t = R_0 + R_1(Z_t - A_t) \quad (24)$$

where R_0 and R_1 are constants with $R_1 > 0$. This calibration procedure tracks down the ratio of adoption expenditures over output whose volatility is matched to that of the data. We assume that the data counterpart for the number of adopted technologies (ie, $A_{t+1} - \phi A_t$), is the number of patent applications.³ The parameter values for the exogenous stochastic process (8)–(9) and (14), are selected in order to approximate the variance, and first-order autocorrelation of the Solow residual and patent applications, together with the correlation between patents and output over medium-term cycles in the data.

³The series considered is the number of patent applications obtained from U.S. Patent and Trademark Office for 1970–2004, and from Historical Statistics of the United States series W–99 for 1948 to 1970.

The amount of debt B_t , is specified as an exogenous process that incorporates a fix component capturing log-term debt and a variable part representing short-term debt. The fix component is a fraction τ^{def} of the unleveraged value of the stock in the steady state q^{UL} . The variable part is specified as a constant fraction τ^{dev} of the factor payments to labor and the intermediate goods. Hence, debt satisfies

$$B_t \equiv \tau^{def} q^{UL} + \tau^{dev} [\gamma(1 - \alpha) (k_t^\alpha l_t^{1-\alpha})^\rho + (1 - \gamma) M_t^\rho] Y_t^{(1-\rho)} \quad (25)$$

We set the parameters $\tau^{def} = 0.2$ and $\tau^{dev} = 0.3$. These values imply that the share of debt over the unleveraged value of the stock in the steady state is 33.05%, and the short-term part of debt is 42.51%, which is in line with different estimates [see Hall (2001), Rajan and Zingales (1995), and Rouwenhorst (1995)]

5 Quantitative Results

This section contains several numerical experiments to assess the model's predictions. Our primary interest is to quantify the importance of several macroeconomic factors on the volatility of stock market values. As a guide to understand the importance of external forces in the dynamics of the model, we first compute impulse-response functions for shocks in TFP, new technologies, and markups. In a second set of experiments we check the ability of the model to reproduce various second-order moments found in the data. We also discuss some other extensions of the model: Taxes, labor market frictions, and monetary policy.

5.1 Impulse-Response Functions

This first set of exercises focuses on the response of the different components of stock market values of Proposition 3.1 to a one-time perturbation of our forcing variables θ , x and ϑ . To gain further intuition we also report the response of output, consumption and the risk-free interest rate. From inspection of all these graphs, it will appear that the model has long propagation effects that may last for over 50 years. This is not actually a major concern in the sequel. As discussed below, some long-term statistics are actually lower than the empirical values.

Figure 7 exhibits the percentage deviation from the steady state value for the stock price q , and its different components, as well as output y and consumption c , and the risk-free interest rate r after an increase in θ by one standard deviation. These effects are a bit stronger and more persistent than in the neoclassical growth model because of the extra propagation mechanism in the market of intermediate goods. An increase in TFP stimulates consumption and capital accumulation. Then, the interest rate goes down to accommodate convergence back to the steady state. Given the functional form for production function (6) positive changes in θ and k raise profits per variety π_t , which jointly with the lower interest rates results in increments in q , V_t^+ , J_t^+ , and ζ_t . Hence, there are further incentives to invest in technology adoption: H_t will go up, and the amount of adopted varieties A_t will increase over time. This last additional channel generates a more persistent response to the shock.

Figure 8 exhibits the response of the above variables to an increase of one standard deviation in Z . Note that the stock market value q and its components go down because the arrival of new technologies will depress the amount of profits π_t per variety. As these technologies need to be adopted, there is now room for investment in technology adoption at the expense of consumption and capital investment. Output peaks later on because adoption of new technologies stirs up the productivity of capital and labor, and so both consumption and investment must rebound. Figure 9 replicates the same experiment for an increase of one standard deviation of the markup ϑ . As expected an increase in ϑ boosts the stock market. Capital investment and consumption go down. (Consumption goes down even in spite of the wealth effect.) There is, however, more activity in the sector of intermediate goods, and hence output does not necessarily slow down.

To summarize, for our benchmark calibration of the model positive changes in θ and ϑ lead to extended increases in the stock price q , and the effects are more pronounced for changes in ϑ . Positive changes in the set of available technologies Z may actually decrease the stock price q , as the arrival of new technologies depresses the price of the existing ones.

5.2 Second-Order Moments

The foregoing exercises furnish us with some clues in testing the ability of the model to replicate observed volatilities and co-movements in both real and financial sectors. We first

focus on our benchmark calibration, and then comment on some other extensions which either have no effect on the volatility of stock values or generate too much volatility in investment, dividends, and earnings. The model simulations are obtained from paths with 3000 observations, where the first 1000 observations have been dropped to avoid initial conditions effects.

To isolate the role of debt, in Table 1 we present volatilities for a model in which the value of debt is equated to zero, so that equity is the only asset in the economy. The volatility of the stock market is 14.98 as compared to 31.41 in the data. The volatilities of all real variables are in line with the business cycle literature. Note that the volatilities of the Solow residual and patents are matched to those of the data as established in our calibration exercise. Observe that the model generates higher volatilities for dividends and earnings. This should be the case because dividends is a policy variable: Companies may want to smooth out dividends (and even earnings) or engage in share repurchases. Hall (2001) concludes that pay-outs to debt holders have been quite erratic, and the pay-out yield (the ratio of total cash extracted by security owners to the market value of equity and debt) has been anything but steady. Moreover, in our model investment is only performed at the firm level – there is no rental capital – and there are no lags or time-to-build for investment.

If adjustment costs are taken out of the model, then the volatility of investment goes from 9.46 percent to 10.59 percent, and the stock market volatility goes from 14.98 percent to 15.29 percent. Therefore, the introduction of adjustment costs leads to a mild drop in the volatility of capital investment. But it actually has a negative effect on the volatility of the stock market, since adjustment costs may crowd out expenditures in technology adoption.

Table 2 refers to the same volatilities for our benchmark calibration of the model with debt. We see that the volatility of the stock market is now 21.52 percent as compared to 31.41 percent in the data. Therefore, the model can provide for over two-thirds of the actual volatility in the data. The introduction of debt then raises the volatility from 14.98 percent to 21.52 percent, which is about a fifty-percent increase. Observe that this increase is attained at the expense of a higher volatility of dividends and earnings, but as already discussed these numbers may still be plausible. From both Tables 1 and 2 we can see that the model lags behind especially regarding the volatility of the stock market for the long-term cyclical component (i.e., in the frequency band of 8 to 50 years), rather than for the shorter term

cyclical component (2 to 8 years). But for the risk-free interest rate the problem is actually inverted: The model generates much less volatility for the shorter term cyclical component. This may suggest some possible effects of active monetary policies and global uncertainty. Table 3 reports on the persistence of the variables in the model and the data. The model is set up to match the persistence of the Solow residual and patents in the data. We observe that the model can reproduce rather well the autocorrelation observed in the data.

Table 4 displays contemporaneous correlations of output with other macro aggregates. In most of the business cycle literature, output in the model is highly correlated with the Solow residual over the shorter term cyclical component, and output is also too highly correlated with the stock market. In our model, we still see the high correlation of output with the Solow residual, but output is only mildly correlated with the stock market. As a matter of fact, our model shows too little correlation of output with the stock market over low frequencies.

Finally, Table 5 contains the contemporaneous correlations of the stock market with the other macro variables. The previous discussion on the correlation of the stock market with output does extend to the components of output. That is, both consumption and investment are not much correlated with the stock market. It is remarkable that the model replicates well the correlation of the stock market with several variables over the long-term cyclical component. Among these variables, we can include Solow residual, patents, and dividends. The model, however, does not match well the correlation of the stock market with the risk-free interest rate. As discussed in Campbell (1999) there is not settled evidence on the correlation of the stock market with interest rates. From Figure 6, it makes more sense to test stock market returns against the risk-free rate – which is well matched by our model.

In summary, the model can generate a sizable part of the volatility of the stock market, and performs fairly well when considering various correlations and co-movements of macro aggregates in the data. As extensions of this analysis, we could test for leads and lags of the stock market over the cycle, and for several correlations on growth rates and returns rather than on levels. A systematic analysis of leads and lags across variables may have to be supported by a more elaborate calibration of the model that allows for a richer dynamic structure for the exogenous stochastic variables.

5.3 Extensions

Taxes : Taxes on corporate profits and dividends can greatly affect the stock market value [McGrattan and Prescott (2005)] as well as endogenous investment and dividends [Hall (2001) and Poterba (2004)]. We have considered an exogenous process for taxes that is meant to fit the evolution of taxes on dividends in the US from data reported in McGrattan and Prescott (2003). This tax policy had a very small effect on the stock market. Moreover, some activist fiscal policies on taxes and allowances for depreciation [Auerbach (2009)] have a damping effect in stock market values. Of course, the effect of taxes may vary depending on whether they are expected or unexpected. Hence, following Poterba (2004) we have analyzed various arbitrary tax policies and performed various experiments. Overall, the outcome of these experiments is that taxes may affect both the volatility of asset values and dividends, but only at the expense of excessive volatility in some real variables such as capital investment and consumption.

Labor Frictions : The model with variable labor does not improve significantly the volatility of the stock market [see Rouwenhorst (1995)]. The introduction of variable labor brings the same problems encountered in the business cycle literature [e.g., Hansen and Wright (1992) and Kydland (1995)]. Namely, in the data hours worked fluctuate more than productivity, and the correlation between hours worked and productivity is close to zero. Moreover, sticky wages, labor market rigidities and additional shocks to the labor markets⁴ seem to have a minor influence in the long-term volatility of the stock market. General and nested CES production functions for capital and labor and intermediate goods did not lead to much improvement of the current results. As a proxy for labor distortions, we have experimented with a persistent shock in the income shares of labor and capital income. This distortion generated too much volatility on capital investment. Figure 10 plots the evolution of the income shares for labor, capital and intermediate goods for our benchmark calibration. We see that the income share of labor hovers around 60 percent and it is also reasonably volatile. This share seems in accord with casual observations, even though proprietors income in the intermediate goods sector can be relocated to either labor or capital.

Monetary Policy : Monetary policy may affect interest rates and hence the behavior of the

⁴These labor frictions were introduced to resolve these puzzles relating hours and productivity.

stock market. Campbell (1999) reviews the international evidence and argues that there is no clear correlation between interest rates and the stock market. Also, the finance literature finds that borrowing constraints and other market frictions have minor effects on the volatility of asset prices [e.g., Heaton and Lucas(1996)]. Here we have run two experiments and in both we have gotten negative results. The first experiment captures the effects of monetary policy by introducing a government bond and public consumption and an exogenous law of motion for the risk-free interest rate. We match the volatility of the exogenous interest rate to that of actual interest rates for the cyclical component at frequencies of 2 to 8 years so that interest rates are more volatile than in our original calibrated model. Briefly, the effects of this policy on the stock market and other macro aggregates are rather minor. In the second experiment we introduce borrowing constraints at both the level of the household and the firm that may limit the ability to consume or invest in a given period. Rather than fixing those constraints to some constant limits as most of the previous literature [e.g., Heaton and Lucas (1996)], we assume that the role of monetary policy is to set up the level of leverage over time [cf. Geanakoplos 2009]. We consider different stochastic policies that exogenously affect the evolution of leverage, and all the results point in the direction that the stock market will remain essentially unaffected; else, capital investment and consumption would fluctuate wildly.

6 Concluding Remarks

This paper explores macroeconomic determinants of asset price volatility in a model with lags in technology adoption under a CES utility function and a mild form of adjustment costs for capital investment. Technologies are embedded in the production of new varieties of intermediate goods. Stocks are impacted by the arrival of new technologies and other shocks affecting the economy. Our analysis centers on an asset pricing equation that decomposes the value of the stock market into the book value of capital, the value of existing technologies, and the option value of adoption and arrival of new technologies.

This general equilibrium setting imposes a lot of discipline in our numerical experiments: Desired levels of volatility of asset prices usually come with pronounced changes in macroeconomic fluctuations. Since the value of the stock market can be computed as the expected

discounted sum of future dividends, the lack of volatility of stock prices is already familiar from the early empirical studies of LeRoy and Porter (1981) and Shiller (1981). But it becomes much harder to deal with in model simulation as one cannot postulate exogenous specifications for the dividend process. To get an idea of related computations, in our simulations of the neoclassical growth model a volatility (i.e. standard deviation) of investment of the order of 6.5 percent translates into a volatility of the capital stock of the order of 2.3 percent. As the volatility of the stock market is about 31.41 percent it is then not surprising that many candidate variables will have a limited role in accounting for observed fluctuations of stock market values.

Thus, we find that taxes, monetary policy interventions, financial and labor market frictions, and various CES formulations of the aggregate production function for labor and capital have minor effects on the long-term volatility of asset market values. Technological innovations can have significant effects if they come along with high markups and TFP changes. Leverage – short-term and long-term debt – seems also quite relevant, but it cannot arbitrarily be increased in the model as it leads to excessive volatility of dividends. Overall, we get that in our model the volatility of the stock market is of the order of 21.52 percent as opposed to 31.41 percent in the data. Therefore, the model can deliver almost 70 percent of the observed volatility. Paradoxically, the model can fully attain the volatility of the cyclical component over a window of frequencies from 2 to 8 years. Therefore, it appears that we may be missing some structural forces or long-term propagation mechanisms affecting the economy and the global markets.

Our model has other advantages over the traditional business cycle literature in which dividends and earnings are not sufficiently volatile⁵ and the risk-free rate has almost zero variability. Our model displays higher volatility of dividends and earnings than in the data,⁶ and the volatility of the risk-free rate is much closer that of the data. Moreover, there are many other implications for financial markets that come from our fundamental asset pricing equation (21). Indeed, in our model the stock market anticipates news and the evolution of

⁵In a closely related exercise that makes full use of our analytical work, Comin, Gertler and Santacreu (2009) claim to get rather low volatility of dividends, and hence we are doubtful that they could account for the volatility of stock values.

⁶As already discussed, companies may want to smooth out dividends; moreover, some further adjustments are needed in reported dividends figures to bring the model closer to the data [Hall (2001)].

exogenous variables, and so stock values are leading indicators which are only mildly correlated with output and other real variables [Campbell and Shiller (1988)]. In fact, fluctuations of price-earnings and price-dividend ratios in our model are quite close to those in the data. Let us briefly mention two extensions of this work. First, our calibration exercise could be improved to account for cross correlations over time. This calibration refinement may actually need to allow for a richer dynamic structure of the exogenous shocks. Hence, the state space may need to be expanded, but it could be effective to obtain higher long-term volatility. Second, the macro finance literature has struggled with several puzzles and empirical regularities, and the present model may be helpful to address some further issues guided by our fundamental asset pricing equation (21).

7 Appendix

Proof: The first order condition (18) and the transversality condition $\lim_{T \rightarrow \infty} E_t \left\{ \frac{\eta_T}{\eta_t} q_T a_T \right\} = 0$ for the household's problem imply

$$q_t = E_t \left\{ \sum_{r=t+1}^{\infty} \frac{\eta_r}{\eta_t} d_r \right\} \quad (26)$$

Similarly, optimization by the firm implies the transversality condition $\lim_{T \rightarrow \infty} E_t \left\{ \frac{\eta_T}{\eta_t} p_t^I k_{T+1} \right\} = 0$ and hence

$$p_t^I k_{t+1} - B_t = E_t \left\{ \sum_{r=t+1}^{\infty} \frac{\eta_r}{\eta_t} d_r^f \right\}. \quad (27)$$

Then, we must show

$$V_t^+ A_t + J_t^+ (Z_t - A_t) + \xi_t = E_t \left\{ \sum_{r=t+1}^{\infty} \frac{\eta_r}{\eta_t} d_r^I \right\}. \quad (28)$$

To this end, we define the following functions

$$V_{t,T}^+ \equiv E_t \left\{ \sum_{r=t+1}^T \frac{\eta_r}{\eta_t} \phi^{r-t} \pi_r \right\} \text{ for } T \geq 1, \text{ and } V_{t,0}^+ = 0. \quad (29)$$

$$\begin{aligned}
J_{t,T}^+ &\equiv E_t \left\{ \sum_{r=1}^T \frac{\eta_{t+r}}{\eta_t} \phi^r \left[\left(\prod_{j=0}^{r-2} (1 - \lambda_{t+j}) \right)^{I_{(r \geq 2)}} \lambda_{t+r-1} V_{t+r,T-r} \right. \right. \\
&\quad \left. \left. - \left(\prod_{j=0}^{r-1} (1 - \lambda_{t+j}) \right) H_{t+s} \right] \right\} \text{ for } T \geq 1, \text{ and } J_{t,0}^+ = 0. \\
\xi_{t,T} &\equiv E_t \left\{ \sum_{r=1}^T \frac{\eta_{t+r}}{\eta_t} (Z_{t+s} - \phi Z_{t+s-1}) J_{t+s,T-s} \right\} \text{ for } T \geq 1.
\end{aligned} \tag{30}$$

It is straightforward to prove convergence of these approximate functions:

$$\lim_{T \rightarrow \infty} V_{t,T}^+ = V_t^+ \tag{31}$$

$$\lim_{T \rightarrow \infty} J_{t,T}^+ = J_t^+ \tag{32}$$

$$\lim_{T \rightarrow \infty} \xi_{t,T}^+ = \xi_t^+ \tag{33}$$

$$V_{t,T}^+ A_t + J_{t,T}^+ (Z_t - A_t) + \xi_{t,T} = E_t \left\{ \sum_{r=t+1}^T \frac{\eta_r}{\eta_t} d_r^I \right\} \tag{34}$$

Then, taking limits in (34) we obtain (28).

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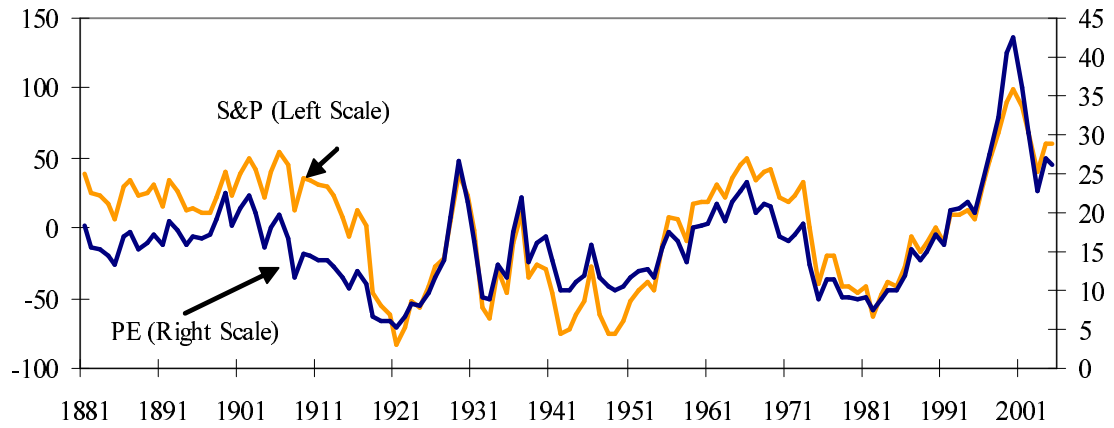
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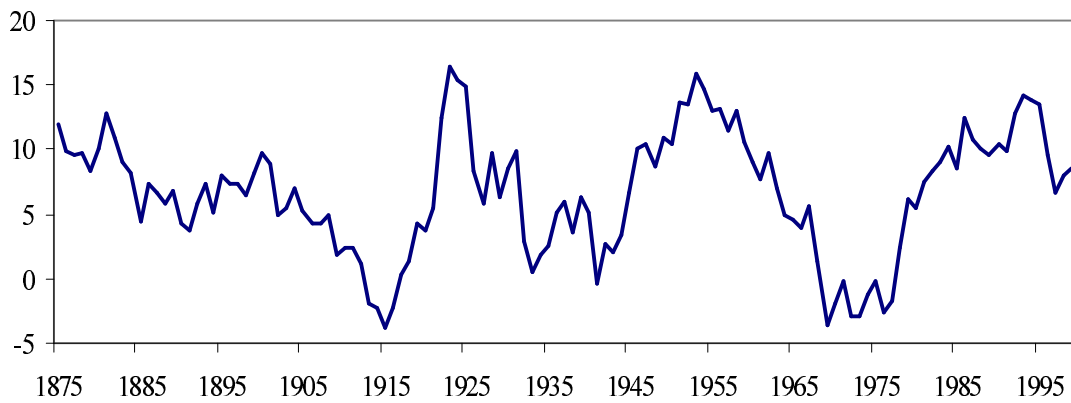
Figure 1: Evolution of S&P and PE



Notes: S&P price index and Price-Earnings ratio. Detrended, annual data from 1881 to 2005.

Source: Robert Shiller's web site: <http://www.econ.yale.edu/shiller/data.htm>.

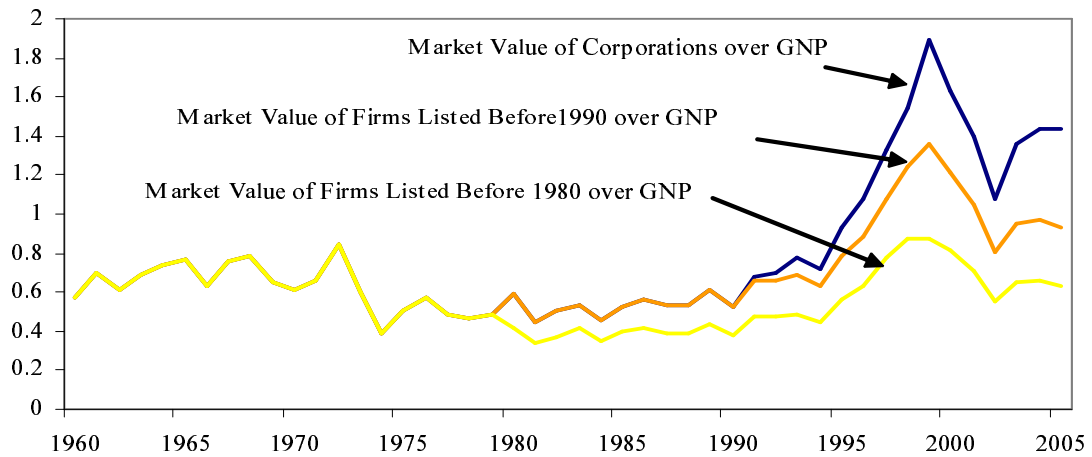
Figure 2: Real Rate of Return on S&P



Notes: Centered ten-year moving average of the continuously compounded log return of S&P index. Annual data from 1881 to 2005.

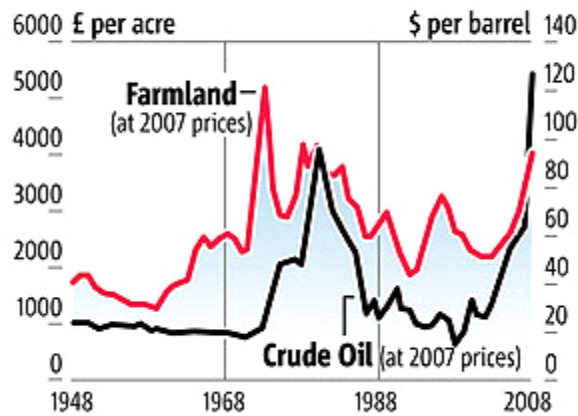
Source: Robert Shiller's web site.

Figure 3: Market Value of Different Vintages



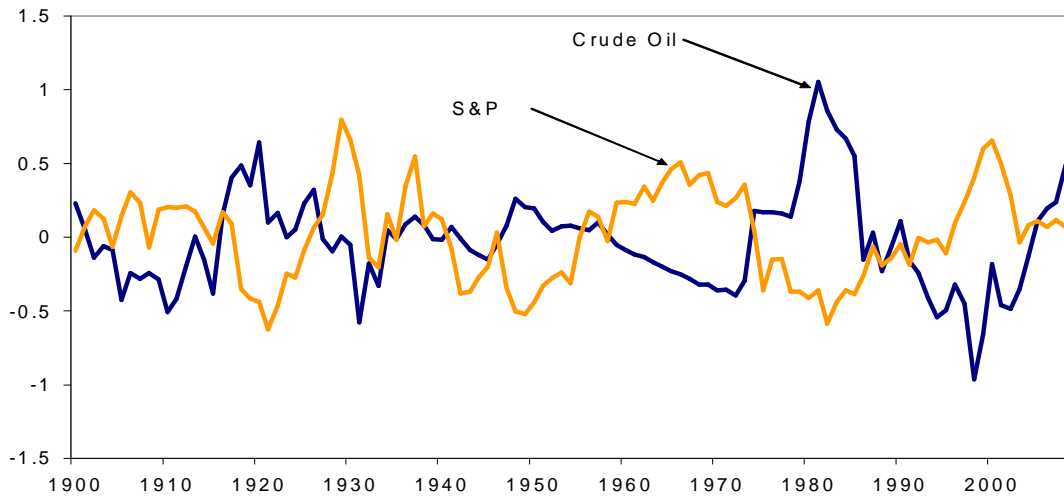
Notes: Market value of corporations over GNP for different vintages. Annual data from 1960 to 2005.
Source: CRSP and NIPA.

Figure 4: Farmland and Crude Oil Prices



Source: <http://www.thisismoney.co.uk>

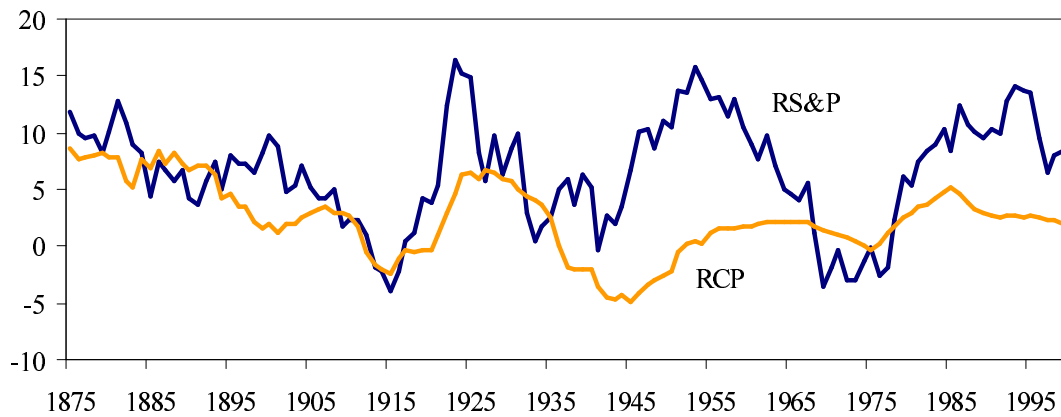
Figure 5: Evolution of S&P and Crude Oil Price



Notes: Real S&P price index and U.S. Real Crude Oil Wellhead Acquisition Price by First Purchasers (Dollars per Barrel). Annual data from 1900-2008. The logarithm of each series have been filtered with a band pass filter that preserves fluctuations for frequencies of 0-50 years.

Source: Robert Shiller's web site and Energy Information Administration.

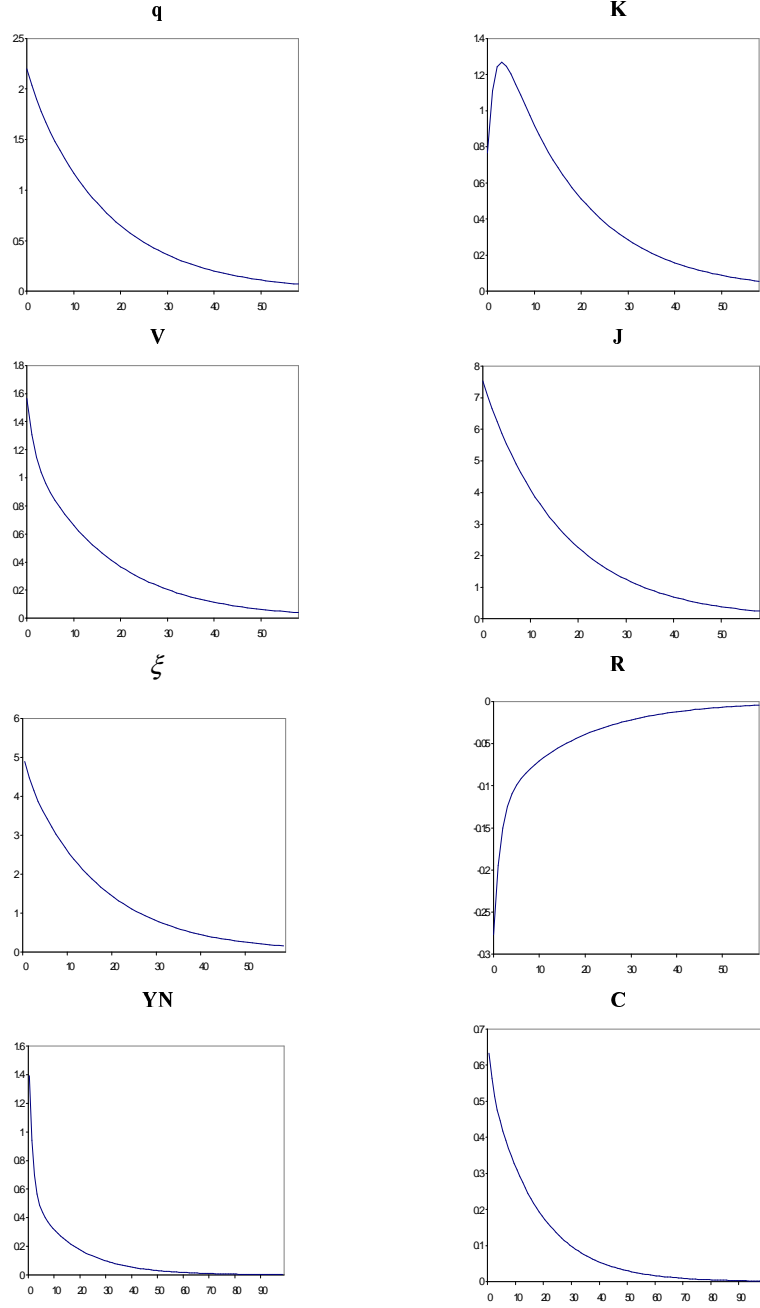
Figure 6: Real Rate of Return on S&P and Short-Term Interest Rate



Notes: Centered ten-year moving average of the continuously compounded log annual return of S&P index and the return on 6-month prime commercial paper. Annual data 1881-2005.

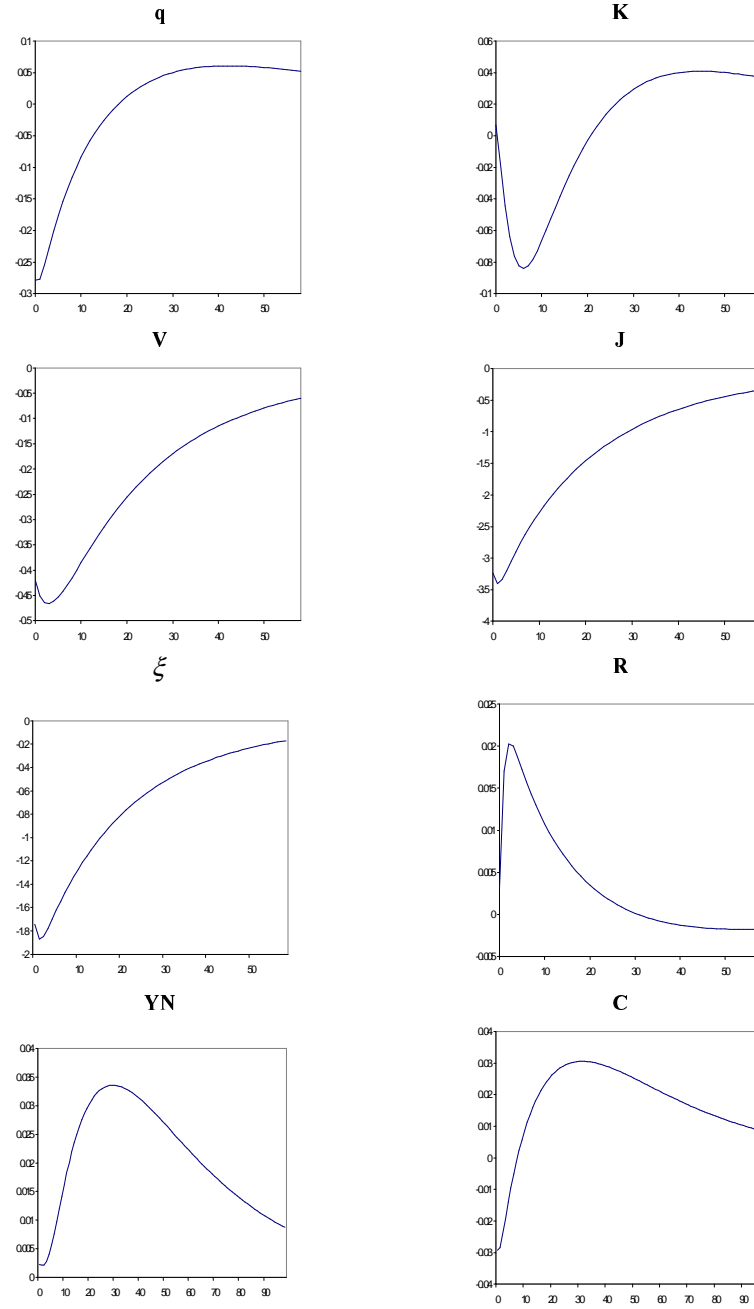
Source: Robert Shiller's web site.

Figure 7: Impulse–Response Functions to a Shock in θ_t



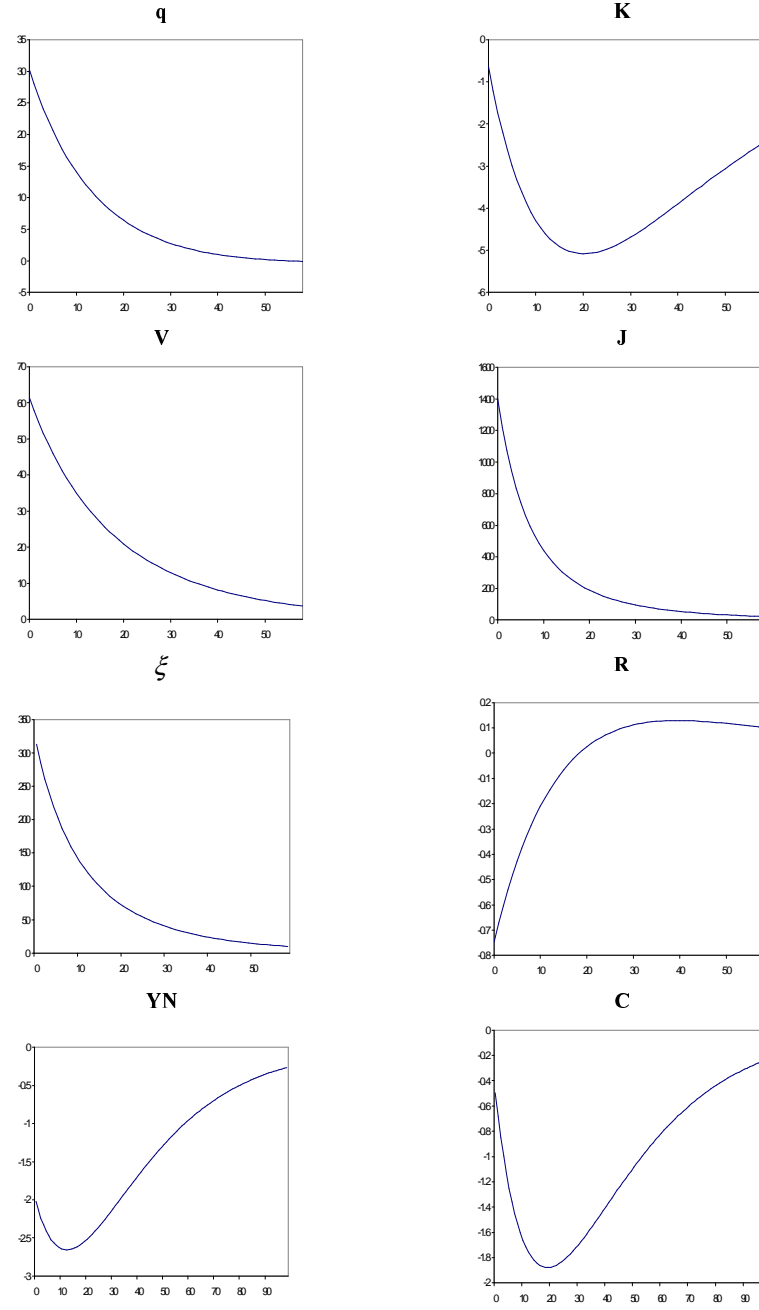
Notes: Response of each variable to a positive, one–standard–deviation shock in θ_t . The variable K denotes capital, V is the value of an installed technology, J is the value of a not–adopted technology, ξ is the present value of technologies available in the future, YN is output, C is consumption, and R is the one–year risk–free interest rate.

Figure 8: Impulse–Response Functions to a Shock in x_t



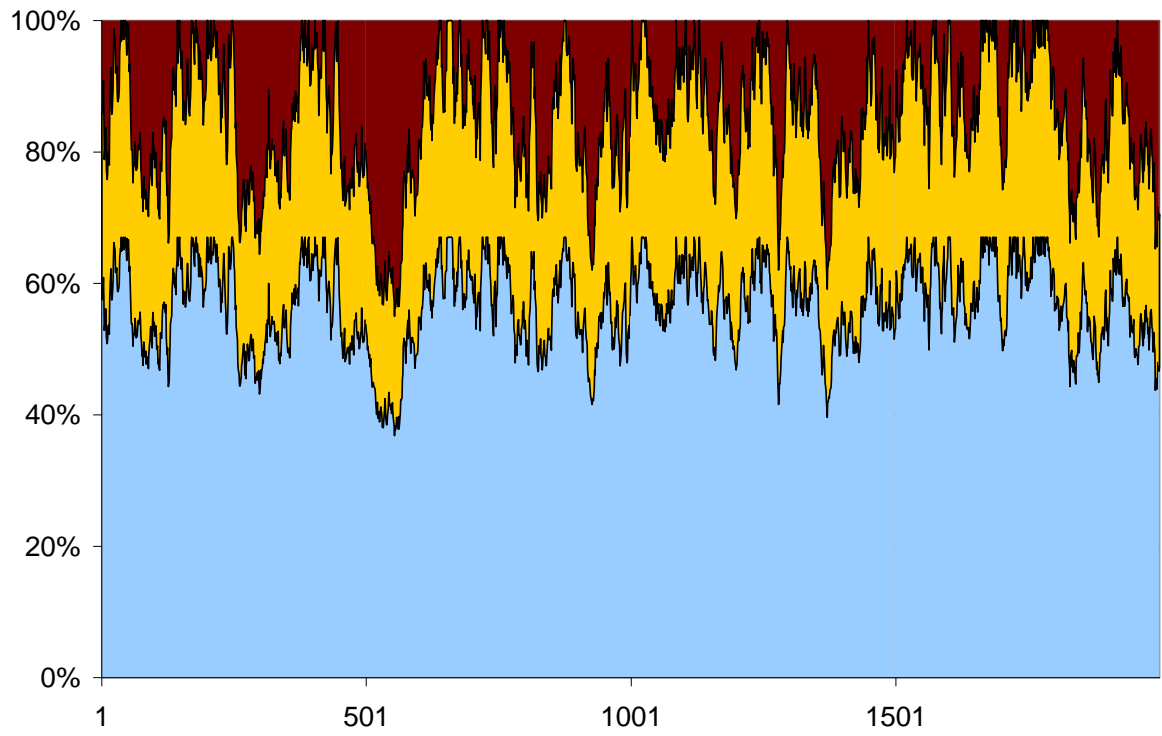
Notes: Response of each variable to a positive, one-standard-deviation shock in x_t .

Figure 9: Impulse-Response Functions to a Shock in ϑ_t



Notes: Response of each variable to a positive, one-standard-deviation shock in ϑ_t .

Figure 10: Factor Shares



Notes: The Blue area is the labor share, the orange/yellow area is the capital share, and the garnet/red area is the intermediate sector's profits over net output.

| Table 1: Standard Deviation | | | | | | |
|-----------------------------|-------------------------|-------|-------------------------|-------|-------------------------|-------|
| | 2-50 | | 2-8 | | 8-50 | |
| | Data | Model | Data | Model | Data | Model |
| Output | 3.46 (2.65; 4.27) | 2.74 | 1.98 (1.53; 2.42) | 1.10 | 2.81 (2.19; 3.42) | 2.51 |
| Consumption | 1.92 (1.38; 2.47) | 1.56 | 0.68 (0.55; 0.81) | 0.39 | 1.79 (1.29; 2.28) | 1.51 |
| Investment | 8.86 (7.09; 10.65) | 9.46 | 4.35 (3.59 ; 5.10) | 4.72 | 7.66 (5.69; 9.63) | 8.19 |
| Solow Residual | 2.43 (1.95; 2.90) | 2.43 | 1.14 (0.85; 1.43) | 1.17 | 2.12 (1.69; 2.55) | 2.12 |
| Patents | 8.03 (6.33; 9.73) | 8.03 | 3.05 (2.26; 3.85) | 3.17 | 7.37 (5.65; 9.09) | 7.42 |
| Stock Market | 31.41 (25.22; 37.58) | 14.98 | 9.28 (7.32; 11.25) | 6.71 | 29.95 (24.08; 35.83) | 13.39 |
| Dividends | 10.27 (8.21; 12.33) | 19.52 | 3.37 (2.01; 4.73) | 7.65 | 9.59 (7.47; 11.71) | 17.96 |
| Earnings | 17.92 (13.45; 22.39) | 18.02 | 11.68 (8.00; 15.36) | 10.49 | 12.53 (10.20; 14.87) | 14.65 |
| Interest Rate | 1.70 (1.18; 2.24) | 1.37 | 1.70 (1.18; 2.24) | 0.74 | 1.84 (1.18; 2.50) | 1.15 |
| Price-Dividends | 32.73 (25.92; 39.54) | 6.71 | 14.20 (11.39; 17.00) | 2.24 | 28.95 (21.25; 36.66) | 6.32 |
| Price-Earnings | 24.65 (19.87; 29.44) | 6.12 | 9.05 (7.12; 10.97) | 4.47 | 22.89 (18.48; 27.30) | 4.18 |

Note: Both series have been filtered for various frequency bands.

| Table 2: Standard Deviation | | | | | | |
|-----------------------------|-------------------------|-------|-------------------------|-------|-------------------------|-------|
| | 2-50 | | 2-8 | | 8-50 | |
| | Data | Model | Data | Model | Data | Model |
| Output | 3.46 (2.65; 4.27) | 2.74 | 1.98 (1.53; 2.42) | 1.10 | 2.81 (2.19; 3.42) | 2.51 |
| Consumption | 1.92 (1.38; 2.47) | 1.56 | 0.68 (0.55; 0.81) | 0.39 | 1.79 (1.29; 2.28) | 1.51 |
| Investment | 8.86 (7.09; 10.65) | 9.46 | 4.35 (3.59 ; 5.10) | 4.72 | 7.66 (5.69; 9.63) | 8.19 |
| Solow Residual | 2.43 (1.95; 2.90) | 2.43 | 1.14 (0.85; 1.43) | 1.17 | 2.12 (1.69; 2.55) | 2.12 |
| Patents | 8.03 (6.33; 9.73) | 8.03 | 3.05 (2.26; 3.85) | 3.17 | 7.37 (5.65; 9.09) | 7.42 |
| Stock Market | 31.41 (25.22; 37.58) | 21.52 | 9.28 (7.32; 11.25) | 9.93 | 29.95 (24.08; 35.83) | 19.08 |
| Dividends | 10.27 (8.21; 12.33) | 38.96 | 3.37 (2.01; 4.73) | 16.01 | 9.59 (7.47; 11.71) | 35.51 |
| Earnings | 17.92 (13.45; 22.39) | 37.50 | 11.68 (8.00; 15.36) | 19.37 | 12.53 (10.20; 14.87) | 32.10 |
| Interest Rate | 1.70 (1.18; 2.24) | 1.37 | 1.70 (1.18; 2.24) | 0.74 | 1.84 (1.18; 2.50) | 1.15 |
| Price-Dividends | 32.73 (25.92; 39.54) | 22.23 | 14.20 (11.39; 17.00) | 10.37 | 28.95 (21.25; 36.66) | 19.67 |
| Price-Earnings | 24.65 (19.87; 29.44) | 20.22 | 9.05 (7.12; 10.97) | 12.06 | 22.89 (18.48; 27.30) | 16.22 |

Note: Both series have been filtered for various frequency bands.

Table 3: Autocorrelation

| | 2-50 | | 2-8 | | 8-50 | |
|-----------------|----------------------|-------|------------------------|-------|----------------------|-------|
| | Data | Model | Data | Model | Data | Model |
| Output | 0.64 (0.47; 0.80) | 0.80 | 0.11 (-0.14; 0.38) | 0.03 | 0.89 (0.77; 1) | 0.95 |
| Consumption | 0.84 (0.68; 0.99) | 0.91 | 0.22 (-0.03; 0.48) | 0.07 | 0.93 (0.80; 1) | 0.97 |
| Investment | 0.72 (0.55; 0.88) | 0.71 | 0.14 (-0.11; 0.40) | 0.01 | 0.91 (0.77; 1) | 0.94 |
| Solow Residual | 0.73 (0.60; 0.87) | 0.73 | 0.10 (-0.16; 0.36) | 0.03 | 0.92 (0.83; 1) | 0.94 |
| Patents | 0.81 (0.70; 0.97) | 0.81 | 0.08 (-0.17; 0.34) | 0.20 | 0.93 (0.88; 1) | 0.92 |
| Stock Market | 0.87 (0.75; 0.98) | 0.74 | -0.02 (-0.28; 0.24) | 0.02 | 0.96 (0.87; 1) | 0.93 |
| Dividends | 0.81 (0.66; 1) | 0.80 | 0.21 (-0.02; 0.44) | 0.12 | 0.94 (0.84; 1) | 0.94 |
| Earnings | 0.51 (0.31; 0.70) | 0.72 | 0.07 (-0.19; 0.34) | 0.14 | 0.86 (0.75; 0.96) | 0.92 |
| Interest Rate | 0.47 (0.19; 0.76) | 0.65 | -0.03 (-0.28; 0.24) | -0.02 | 0.90 (0.75; 1) | 0.93 |
| Price-Dividends | 0.76 (0.62; 0.90) | 0.74 | -0.07 (-0.26; 0.11) | 0.01 | 0.95 (0.85; 1) | 0.94 |
| Price-Earnings | 0.82 (0.67; 0.96) | 0.64 | -0.09 (-0.20; 0.01) | 0.14 | 0.95 (0.86; 1) | 0.91 |

Note: Both series have been filtered with a band pass filter for various frequency bands.

Table 4: Correlation with Output

| | 2-50 | | 2-8 | | 8-50 | |
|-----------------|------------------------|-------|-------------------------|-------|------------------------|-------|
| | Data | Model | Data | Model | Data | Model |
| Consumption | 0.74 (0.49; 0.98) | 0.92 | 0.83 (0.70; 0.98) | 0.86 | 0.74 (0.42; 1) | 0.95 |
| Investment | 0.58 (0.22; 0.94) | 0.94 | 0.83 (0.68; 0.97) | 0.96 | 0.50 (0.04; 0.95) | 0.95 |
| Solow Residual | 0.62 (0.33; 0.89) | 0.92 | 0.83 (0.69; 0.98) | 0.99 | 0.55 (0.18; 0.89) | 0.90 |
| Patents | 0.31 (0.03; 0.58) | 0.00 | -0.22 (-0.48; 0.02) | -0.02 | 0.49 (0.23; 0.76) | 0.00 |
| Stock Market | 0.47 (0.17; 0.73) | 0.17 | -0.04 (-0.30; 0.21) | 0.61 | 0.61 (0.30; 0.93) | 0.07 |
| Dividends | 0.57 (0.33; 0.91) | -0.09 | 0.20 (-0.05; 0.45) | 0.50 | 0.66 (0.41; 0.93) | -0.21 |
| Earnings | 0.40 (0.22; 0.61) | 0.11 | 0.45 (0.22; 0.68) | 0.61 | 0.48 (0.20; 0.72) | -0.01 |
| Interest Rate | -0.26 (-0.63; 0.11) | -0.16 | -0.50 (-0.73; -0.28) | -0.67 | -0.11 (-0.56; 0.34) | -0.02 |
| Price-Dividends | 0.22 (-0.05; 0.50) | 0.33 | -0.40 (-0.58; -0.22) | -0.19 | 0.43 (0.09; 0.77) | 0.45 |
| Price-Earnings | 0.37 (0.05; 0.70) | -0.03 | -0.11 (-0.30; 0.06) | -0.47 | 0.52 (0.17; 0.87) | 0.11 |

Notes: Both series have been filtered with a band pass filter for various frequency bands.

Table 5: Correlation with the Stock Market

| | 2-50 | | 2-8 | | 8-50 | |
|-----------------|-----------------------|-------|------------------------|-------|-----------------------|-------|
| | Data | Model | Data | Model | Data | Model |
| Output | 0.47 (0.21; 0.72) | 0.17 | -0.04 (-0.29; 0.21) | 0.61 | 0.61 (0.26; 0.94) | 0.07 |
| Consumption | 0.61 (0.25; 0.95) | 0.24 | 0.10 (-0.13; 0.34) | 0.61 | 0.68 (0.29; 1) | 0.20 |
| Investment | 0.14 (-0.30; 0.58) | 0.02 | -0.21 (-0.46; 0.04) | 0.51 | 0.21 (-0.27 0.67) | -0.11 |
| Solow Residual | 0.09 (-0.35; 0.50) | 0.03 | 0.30 (0.11; 0.87) | 0.58 | 0.05 (-0.46; 0.55) | -0.12 |
| Patents | 0.53 (0.27; 0.78) | 0.62 | -0.07 (-0.33; 0.18) | 0.12 | 0.63 (0.34; 0.89) | 0.73 |
| Dividends | 0.75 (0.53; 0.97) | 0.88 | 0.25 (0; 0.50) | 0.77 | 0.81 (0.58; 1) | 0.91 |
| Earnings | 0.19 (-0.08; 0.48) | 0.90 | 0.09 (-0.16; 0.35) | 0.85 | 0.28 (-0.16; 0.70) | 0.92 |
| Interest Rate | 0.18 (-0.24; 0.61) | -0.75 | -0.06 (-0.26; 0.15) | -0.72 | 0.28 (-0.25; 0.80) | -0.76 |
| Price-Dividends | 0.85 (0.69; 1) | -0.58 | 0.90 (0.39; 1) | -0.24 | 0.91 (0.72; 1) | -0.67 |
| Price-Earnings | 0.95 (0.86; 1) | -0.61 | 0.59 (0.53; 0.65) | -0.54 | 0.96 (0.87; 1) | -0.65 |

Note: Both series have been filtered with a band pass filter for various frequency bands.