INFLATION AND GROWTH FORECAST ERRORS AND THE SACRIFICE RATIO OF MONETARY POLICY IN THE EURO AREA

2025

BANCODE **ESPAÑA**

Eurosistema

Documentos de Trabajo N.º 2516

Corinna Ghirelli, Javier J. Pérez and Daniel Santabárbara

INFLATION AND GROWTH FORECAST ERRORS AND THE SACRIFICE RATIO OF MONETARY POLICY IN THE EURO AREA

INFLATION AND GROWTH FORECAST ERRORS AND THE SACRIFICE RATIO OF MONETARY POLICY IN THE EURO AREA (*)

Corinna Ghirelli (**)

BANCO DE ESPAÑA

Javier J. Pérez^(***)

BANCO DE ESPAÑA

Daniel Santabárbara (****)

BANCO DE ESPAÑA

(*) We are grateful to Evi Pappa, Luis Julián Álvarez, Pablo Burriel, Gergely Ganics, Florens Odendahl, Carlos Thomas and an anonymous referee as well as participants of the Banco de España seminar for their useful comments. We also thank Alba Villahermosa, Fructuoso Borrallo, Paula Nieto and Juan Carlos Caballero for their valuable help and assistance.

(**) e-mail: corinna.ghirelli@bde.es (***) e-mail: javierperez@bde.es (****) e-mail: daniel.santabarbara@bde.es

Documentos de Trabajo. N.º 2516

March 2025

The Working Paper Series seeks to disseminate original research in economics and finance. All papers have been anonymously refereed. By publishing these papers, the Banco de España aims to contribute to economic analysis and, in particular, to knowledge of the Spanish economy and its international environment.

The opinions and analyses in the Working Paper Series are the responsibility of the authors and, therefore, do not necessarily coincide with those of the Banco de España or the Eurosystem.

The Banco de España disseminates its main reports and most of its publications via the internet at the following website: http://www.bde.es.

Reproduction for educational and non-commercial purposes is permitted provided that the source is acknowledged.

© BANCO DE ESPAÑA, Madrid, 2025

ISSN: 1579-8666 (on line)

Abstract

This paper investigates the relationship between inflation and GDP growth forecast errors and the expected monetary policy stance in the euro area during the monetary policy cycle of 2022-2024, when inflation was well above the ECB's target. Under rational expectations, forecasts of monetary contractions should be unrelated to subsequent inflation and growth forecast errors. On the contrary, we find that expected monetary policy tightening has been associated with higher than projected GDP growth, suggesting a lower monetary policy effect than that factored in by (ECB/Eurosystem and IMF) forecasters. In other words, forecasters overestimated the monetary multiplier. At the same time, monetary policy tightening has been associated with lower than expected inflation, suggesting an underestimation of the monetary multiplier on inflation. Putting these two stylized facts together implies that forecasters overestimated the sacrifice ratio during the last monetary policy tightening cycle. Our findings suggest that forecasters may have inaccurately perceived the recent inflationary crisis in the euro area as predominantly supply-driven, underestimating its demand-driven component. This led to the belief that monetary policy in the euro area would be exceedingly costly in terms of output.

Keywords: forecast errors, monetary policy multipliers, sacrifice ratio.

JEL classification: C53, E27, E62, E52, E58.

Resumen

Este trabajo investiga la relación entre los errores de previsión de la inflación y el crecimiento del PIB, y el tono esperado de la política monetaria en el área del euro durante el ciclo de política monetaria de 2022-2024, cuando la inflación estaba muy por encima del objetivo del Banco Central Europeo (BCE). Bajo expectativas racionales, las previsiones de contracciones monetarias deberían ser independientes de los errores de previsión de la inflación y del crecimiento subsiguientes. Sin embargo, encontramos que las previsiones de endurecimiento de la política monetaria están asociadas con un crecimiento del PIB superior al anticipado, lo que sugiere un menor efecto de la política monetaria en la actividad que el incorporado por los expertos del BCE/Eurosistema y del FMI. En otras palabras, estas proyecciones sobreestimaron la repercusión del multiplicador de la política monetaria sobre el crecimiento. Al mismo tiempo, las previsiones de endurecimiento de la política monetaria están relacionadas con una inflación menor que la esperada, lo que sugiere la subestimación del efecto del multiplicador de la política monetaria sobre la inflación. La combinación de estos dos hechos estilizados implica que estos expertos sobreestimaron la ratio de sacrificio durante el último ciclo de endurecimiento de la política monetaria. Esta evidencia sugiere que se habría percibido la reciente crisis inflacionaria en el área del euro, de manera inexacta, como predominantemente impulsada por factores de oferta, lo que llevó a subestimar su componente de demanda y a la creencia de que la política monetaria en el área del euro sería excesivamente costosa en términos de actividad.

Palabras clave: errores de previsión, multiplicadores de política monetaria, ratio de sacrificio.

Códigos JEL: C53, E27, E62, E52, E58.

1 Introduction

This paper investigates the relationship between inflation and GDP growth forecast errors and the expected monetary policy stance in the euro area during the monetary policy cycle of 2022-2024, when inflation was well above the target of the ECB (Lane, 2024). Under rational expectations, forecasts of monetary contractions should be unrelated to subsequent inflation and growth forecast errors. On the contrary, we find that forecasters (ECB/Eurosystem's and IMF's) overestimated the impact of monetary policy on growth, given that we find a positive relation between underestimations of monetary policy restrictiveness and subsequent growth forecast errors. At the same time, we find that such underestimation is negatively associated with subsequent inflation forecast errors. This implies that, during the last monetary policy tightening cycle, forecasters overestimated the sacrifice ratio (Ball, 1994), which in monetary policy represents the cost of output lost per 1% reduction in the inflation rate.¹

To reach these conclusions, we draw on the approach of Blanchard and Leigh (2013) and Blanchard and Leigh (2014), who addressed a related question regarding the fiscal consolidation programs planned for 2010-2011. They constructed a database of GDP forecast errors from IMF projections and regressed these against the planned fiscal stance, measured by changes in the structural fiscal balance as a percentage of potential GDP. Their study focused on European economies and found that the IMF forecasts underestimated the recessive impacts of austerity, revealing a negative correlation between the fiscal consolidation forecasts and subsequent growth prediction errors. Similarly to the 2010-2011 period, where European economies faced synchronized fiscal consolidation post-financial crisis, the 2022-2023 period saw central banks globally implementing tightening monetary measures. While Blanchard and Leigh (2013) focused on planned fiscal policies, we examine expected monetary policies, using planned fiscal policy as a control variable.² We focus on euro area countries and use two datasets: ECB/Eurosystem Staff projections and IMF WEO forecasts.

Blanchard and Leigh's approach has been followed by several recent studies. Brinca et al. (2021) extend Blanchard and Leigh (2013, 2014) by using the same IMF data and methodology, and adding income inequality measures from the European Union Statistics on Income and Living Conditions (EU-SILC) dataset. They find that fiscal consolidation in 2010-2011 exacerbated recessionary effects due to inequality in Europe. Similarly, Fatás

as exogenous assumptions for future interest rates.

¹In other words, how much GDP growth one has to forego to reduce inflation through monetary tightening.

²In central bank forecasting exercises, future fiscal policies are typically based on government plans, whereas future monetary policies are usually proxied by financial market expectations (futures), incorporated

and Summers (2018) expand Blanchard and Leigh (2013)'s methodology to broader horizons and potential output, examining fiscal consolidations during 2009-2011. They find that these consolidations had prolonged negative impacts on both actual and potential GDP, suggesting that austerity's costs extend beyond immediate GDP effects. Lewis and Pain (2014) conduct a similar analysis using OECD projections for 2007-2012. They find weak evidence that stronger projected fiscal consolidations are associated with larger GDP growth errors, particularly when Greece is included. Suyker and Mohlmann (2015) update Blanchard and Leigh (2013)'s work with data up to 2015. They find a significant negative coefficient for 2009-2011, but no such relationship for 2012-2015, suggesting that IMF forecasters correctly incorporated the impact of fiscal policy during the sovereign debt crisis but underestimated austerity costs during the financial crisis.

Few recent studies have also explored related themes. Darvas (2018) discusses significant forecasting errors by the European Central Bank (ECB) over the past five years, suggesting that shifts in behavioral relationships within ECB models may have influenced policy decisions. Granziera et al. (2024) investigates systematic forecast errors in the ECB's inflation projections and the implications these errors have for the credibility of the bank's monetary policy. On a similar note, Chahad et al. (2024) investigates ECB forecast errors in recent years, a period marked by significant challenges in predicting inflation due to large, unprecedented shocks (the COVID-19 pandemic and Russia's war against Ukraine) impacting the euro area economy. The paper delves into potential explanations for these forecast errors, analyzing the dominant role of global factors, the evolving impact of energy prices, and other exceptional shocks that contributed to the inaccuracies in inflation projections.

Our paper makes three contributions to the literature. First, it is the first study to assess how effectively the most recent cycle of monetary tightening has been incorporated into forecasts during the recent tightening period. Second, our findings contribute to the ongoing policy discussion by showing that in the euro area, expected monetary policy tightening has been associated with higher GDP growth and lower inflation than projected. This suggests that forecasters overestimated the sacrifice ratio—the cost of reducing inflation through monetary policy. Third, our paper adds to the existing literature on ECB's forecast errors by linking the errors during the monetary policy tightening period with the perceived monetary policy stance, particularly in terms of assessing the actual conduct of monetary policy during this extraordinary period. These insights not only clarify recent forecast errors but also provide valuable guidance for future monetary policy decisions.

The rest of the paper is structured as follows. In Section 2 we describe the data. Section 3 outlines the empirical framework, and in Section 4 we present our results. Finally, Section 5 outlines a number of robustness exercises, while Section 6 concludes with final observations and remarks.

2 Data

We analyze data from the ECB/Eurosystem quarterly projection database, covering 19 euro area countries³ over 9 projection exercises from 2022Q1 to 2024Q1. Each vintage uses the database from that quarter's forecast for a specific country.⁴ For robustness, we also use a similar database from the IMF biannual forecasts⁵ over the same period, covering 4 biannual projection exercises (April 2022 to October 2023) for 27 advanced economies, including euro area countries.⁶

Our baseline model focuses on yearly GDP growth rate, headline inflation, and monetary policy stance. Forecast errors are defined as the difference between observed and forecasted values at a specific horizon. Given the medium-term target of monetary policy (two years ahead), our baseline model uses a two-year forecast horizon: h = 2. For observed GDP growth data, we use the first estimate available for forecast projections to ensure comparability across vintages. If GDP data are not yet observed, we use the latest forecast available at the time of writing, i.e., the 2024Q2 forecast exercise for ECB/Eurosystem forecasts and the April 2024 WEO for IMF forecasts.

Measuring the monetary policy stance presents several challenges due to the existence of different instruments (e.g., interest rates, balance sheet policies, central bank communication) and their lagged effects. Non-linear effects and thresholds further complicate assessment. The most common measure is the natural rate of interest gap, which is the difference between the short-term real interest rate and the natural rate of interest (r-star). The natural rate is the short-term real interest rate that corresponds to a neutral monetary policy stance (Taylor, 1993). If the actual interest rate is below r-star, the policy is expansionary; if above, it is contractionary. Ideally, we would use the gap between (expected) short-term

 $^{^{3}}$ We exclude Croatia (HR) as it joined the EMU in 2023, which would unbalance the projection panel.

⁴Country-specific Eurosystem projections are published on the respective National Central Bank websites, while euro area aggregate projections are on the ECB website.

⁵IMF forecasts are in the World Economic Outlook (WEO) and published online.

⁶The EMU countries are: Austria, Belgium, Cyprus, Estonia, Finland, France, Germany, Greece, Ireland, Italy, Latvia, Lithuania, Luxembourg, Malta, Netherlands, Portugal, Slovakia, Slovenia, Spain (19). The extra-EMU countries are: Australia, Canada, Japan, New Zealand, Sweden, Switzerland, United Kingdom, United States (8).

real rates and r-star for each economy. However, r-star is a theoretical and unobservable concept, requiring estimation that is model-dependent.⁷ The estimation of r-star is debated extensively (e.g., Buncic 2024; Benigno et al. 2024; Fiorentini et al. 2018).⁸ Measuring the natural interest rate is also difficult in real-time (Forbes et al., 2024). Some practitioners prefer not to use r-star estimates for these reasons. Alternatively, Forbes et al. (2024) use the nominal interest rate as a direct measure of the monetary policy stance, arguing it avoids the substantial measurement error of more complex measures.

Our analysis focuses on euro area countries, which should share a common natural rate of interest. Given its persistence, we approximate the natural rate of interest gap using the real short-term interest rate, with r-star absorbed by the constant. We use projected inflation to obtain a measure of projected real interest rates as our measure of monetary stance. We also perform robustness checks using alternative measures: the nominal short-term interest rate as in Forbes et al. (2024), and the long-term real interest rate, calculated for the euro area considering the central bank's mid-term inflation objective of 2%.

We include additional controls for robustness, particularly relevant post-COVID and during the energy crisis. First, we consider the fiscal policy stance, measured as in Blanchard and Leigh (2013) by the change in the general government structural fiscal balance as a percent of potential GDP, estimated in each vintage.¹¹ Second, given the inflationary crisis and the missed dynamics of the energy component of CPI inflation, we consider forecast errors in energy commodity prices. Energy prices are expressed as a weighted average of gas (42%) and oil (58%), consistent with the Eurosystem projections. Oil prices refer to Brent crude oil spot and futures prices, while gas prices refer to the Dutch TTF gas spot and futures prices. These are the series used in the Eurosystem projections.

 $^{^{7}}$ Lubik et al. (2024) provide estimates of r-star for several countries, including the euro area, using the same methodology as the Richmond Fed's estimate for the US. Other estimates can be found here.

⁸Benigno et al. (2024) state that "The measurement of *r-star* is surrounded by very high uncertainty, making it a blurry guidepost for monetary policy".

⁹In Eurosystem projections, market expectations of interest rates at the forecast exercise cutoff-date are used as proxies for future short-term interest rates. We make the same assumption. The WEO database does not provide interest rate forecasts, so we construct our measure of expected monetary policy stance using external data. We gather future short-term interest rates expected by financial markets at the IMF forecast cutoff-date for each country in the WEO database. We include extra-EMU advanced economies for which short-term interest rate futures can be collected. Further details are in Section A of the Supplementary Appendix.

¹⁰Using long-term interest rates as a measure for monetary policy stance relies on the idea that they reflect agents' perceptions of the future path of short-term interest rates. A rise in long-term rates suggests a tightening of monetary policy expectations. However, long-term rates also include risk premia associated with uncertain economic conditions and interest rates.

¹¹For example, in the 2022Q1 exercise, the fiscal stance is the difference between the structural fiscal balance estimated in March 2022 for 2022 and for 2021.

3 Methodology

Our research examines whether GDP growth and inflation forecast errors are systematically associated with monetary policy stance forecasts. Following Blanchard and Leigh (2014), we regress two-year-ahead forecast errors for inflation and real GDP growth (2022-2024) against expected monetary policy stance. Under rational expectations and correct forecasting models, the coefficient for expected monetary policy stance should be zero. A significant non-zero coefficient suggests a bias in estimating the impact of monetary policy on the real economy and inflation. In the case of growth forecasts, a negative β indicates underestimation of the monetary multiplier, while a positive β suggests overestimation.

Consider the ECB/Eurosystem quarterly forecasts for GDP growth two years ahead. Each vintage, v, corresponds to 2022Q1-2024Q1, with t representing the years 2022, 2023, and 2024. The estimated equation is therefore

$$\Delta Y_{i,t+2} - \Delta Y_{i,t:t+2|v}^F = \alpha + \lambda_q + \beta_R R_{i,t:t+2|v}^F + {}_{i,t:t+2}$$
(1)

The dependent variable is the forecast error for annual GDP growth (ΔY) in year t + 2. The annual GDP growth forecast is $\Delta Y_{i,t:t+2|v}^F = f\Delta Y_{t+2,i}|\Omega_v$, where f is the forecast conditional on Ω_v , the information set available early in each projection exercise v. $R_{i,t:t+2|v}^F = fR_{t+2,i}|\Omega_v$ represents the forecast of the short-term interest rate in t+2 ($R_{i,t+2}$), where the forecast is again conditional on on Ω_v . In addition, we control for energy commodity prices forecast errors, which allow to account for the inflationary crisis and the missed dynamics of the energy component of CPI inflation.¹² Finally, the specification includes quarter fixed effects, and we correct standard errors for serial correlation of type MA(2) using the Newey-West procedure.¹³

Quarterly forecast exercises target yearly variables (GDP growth and inflation), making forecasts within the same calendar year non-comparable due to changing information sets and decreasing uncertainty. This is known as *fixed-event forecast*. To improve comparability, we follow Dovern et al. (2012) and shift to a fixed-horizon framework, transforming data into an average of forecasts for the current and next two calendar years (weighted by their share

¹²In the regression of forecast errors for inflation, the inclusion of forecast errors for energy prices may suffice as a global control variable. However, for the GDP growth forecast errors, we carry out a robustness check which includes euro area GDP growth forecast error as additional control variable. This would help account for potential endogeneity between forecast errors for individual countries and the monetary policy stance. Results are reported in Table C.14 of the Online Appendix.

¹³To test the robustness of the standard error estimates, we run the analysis experimenting with a higher lag. Results show, as expected, slightly bigger standard error, and strongly confirm our baseline results. This robustness analysis is available upon request.

in the forecasting horizon), serving as a proxy for the fixed horizon forecast for the next 8 quarters (two years). For more details, see Section B of the Appendix. The fixed-horizon setting ensures that different projections of the same calendar year share the same horizon, improving comparability. However, the information set changes throughout the year. To account for this potential bias, we control for quarter of origin dummies in the fixed-horizon setting. The discussion above is easily transferable to IMF bi-annual WEO forecasts. The next section presents and discusses the main findings.

Table 1: Annual GDP forecast errors on monetary policy stance

	F	ECB/ESC	B Forecas	sts		IMF WEO) Forecast	ts
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real R [†]	0.29***	0.28***	0.31***	0.31***	0.40***	0.47***	0.45***	0.53***
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
$\mathrm{Fiscal}^{\ddagger}$		-0.14	-0.13	-0.13		0.55**	0.56**	0.55*
		(0.41)	(0.44)	(0.47)		(0.05)	(0.05)	(0.05)
Energy [§]			-0.01	-0.01			0.00	-0.00
			(0.29)	(0.41)			(0.49)	(0.94)
q = 2				0.07				
				(0.67)				
q = 3				-0.00				
				(0.99)				
q = 4				0.08				
				(0.73)				
half=2								-0.37
								(0.32)
Constant	-0.27**	-0.23**	-0.27**	-0.31**	-0.32**	-0.46***	-0.36*	-0.23
	(0.02)	(0.03)	(0.01)	(0.04)	(0.02)	(0.00)	(0.06)	(0.37)
Obs	171	171	171	171	76	76	76	76
R2	0.20	0.20	0.21	0.21	0.20	0.28	0.28	0.29

Note: Robust p-values in parentheses. The dependent variable is the annual GDP forecast error (observed minus forecasted). Two-year horizon considered. Data transformed to fixed-horizon setting. Sample: 19 euro area countries.

 $^{^{\}dagger}$ Real R refers to real short term interest rate forecast.

[‡] Fiscal refers to fiscal stance forecast.

[§] Energy refers to energy price forecast errors.

^{***}p < 0.01, **p < 0.05, *p < 0.1.

¹⁴Specific features may be associated with projection exercises of specific quarters. For instance, first and third quarter forecasts are conducted by the ECB, while others are jointly realized by all Eurosystem central banks.

4 Results

Table 1 shows how annual growth forecast errors are associated with monetary policy forecasts. The table has two panels: one with ECB/Eurosystem data and another with IMF WEO data. The first column in each panel reports the baseline specification, regressing only monetary policy stance on GDP growth. Subsequent columns add relevant controls one at a time.

We control for planned fiscal policy stance (columns (2) and (6)), as governments have taken anti-inflationary measures and withdrawn some COVID support initiatives. We also control for forecast errors in energy prices due to their significant role during the inflationary period. The last column includes quarter of origin dummies to account for differences across projection exercises (ECB/Eurosystem dataset) of half-yearly dummies (IMF WEO dataset).

Table 2: Annual inflation forecast errors on monetary policy stance

	E	CB/ESC	B Forecas	\mathbf{ts}]	MF WEO	Forecast	S
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real R^{\dagger}	-0.32***	-0.30***	-0.46***	-0.42***	-0.53***	-0.47***	-0.46***	-0.11
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.56)
$Fiscal^{\ddagger}$		0.37	0.32	0.28		0.49**	0.48**	0.40**
		(0.12)	(0.16)	(0.20)		(0.03)	(0.03)	(0.04)
Energy [§]			0.03***	0.03***			-0.00	-0.01**
			(0.00)	(0.00)			(0.86)	(0.01)
q = 2				-0.65***				
				(0.00)				
q = 3				-0.62**				
				(0.01)				
q = 4				-0.96***				
				(0.00)				
half=2								-1.66***
								(0.00)
Constant	-0.07	-0.17*	0.10	0.55***	-0.04	-0.16	-0.19	0.38*
	(0.39)	(0.09)	(0.36)	(0.00)	(0.71)	(0.17)	(0.34)	(0.07)
Obs	171	171	171	171	76	76	76	76
R2	0.17	0.20	0.37	0.42	0.24	0.29	0.29	0.41

Note: Robust p-values in parentheses. The dependent variable is the annual inflation forecast error (observed minus forecasted). Two-year horizon considered. Data transformed to fixed-horizon setting. Sample: 19 euro area countries.

[†] Real R refers to real short term interest rate forecast.

 $^{^{\}ddagger}$ Fiscal refers to fiscal stance for ecast.

[§] Energy refers to energy price forecast errors.

^{***}p < 0.01, **p < 0.05, *p < 0.1.

Table 3: IMF WEO forecast errors on monetary policy stance

		GDP fore	ecast erro	r	Ir	ıflation fo	recast err	or
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real R^{\dagger}	0.01	-0.00	-0.00	-0.00	0.07***	0.03	0.03	0.05**
	(0.25)	(0.85)	(0.97)	(0.96)	(0.00)	(0.11)	(0.14)	(0.04)
$RR \times EA^{\P}$	0.39***	0.42***	0.41***	0.40***	-0.56***	-0.49***	-0.49***	-0.23
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.14)
$\mathrm{Fiscal}^{\ddagger}$		0.23	0.25	0.25		0.46***	0.45***	0.28*
		(0.24)	(0.22)	(0.23)		(0.01)	(0.01)	(0.05)
Energy [§]			0.00	0.00			-0.00	-0.01***
			(0.40)	(0.41)			(0.90)	(0.00)
half=2				0.05				-1.46***
				(0.85)				(0.00)
Constant	-0.30***	-0.39***	-0.30**	-0.31	0.08	-0.11	-0.13	0.36**
	(0.00)	(0.01)	(0.05)	(0.10)	(0.45)	(0.27)	(0.42)	(0.03)
Obs	108	108	108	108	108	108	108	108
R2	0.20	0.22	0.23	0.23	0.21	0.26	0.26	0.41

Note: Robust p-values in parentheses. The dependent variable is the annual inflation and GDP growth forecast error (observed minus forecasted). Two-year horizon considered. Data transformed to fixed-horizon setting. Sample: advanced economies with available short-term interest rate futures (19 EMU countries and 8 extra-EMU countries: AUS, CAN, JPN, NZL, SWE, CHE, GBR, USA).

The baseline analysis focuses on two-year-ahead forecasts, using data transformed to the fixed-horizon structure.

In all specifications, the coefficients for the monetary policy stance are positive ($\beta_R > 0$) and significant, indicating a positive association between monetary policy forecast and growth forecast errors. This suggests that the growth effect of monetary policy has been milder than forecasters assumed. This result is robust across specifications and forecasters, holding for both ECB/Eurosystem and IMF projections.

Table 2 has a similar structure but considers inflation forecast error as the dependent variable. Here, the coefficients for monetary policy stance errors are negative ($\beta_R < 0$) and significant, indicating a negative association between monetary policy forecast and inflation forecast errors. This suggests that the effect of monetary policy has been stronger than forecasters assumed. This finding is also robust across model specifications and datasets of forecasts.

 $^{^{\}dagger}$ Real R refers to real short term interest rate forecast.

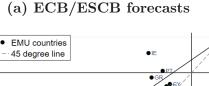
 $^{^{\}ddagger}$ Fiscal refers to fiscal stance forecast.

 $[\]S$ Energy refers to energy price forecast errors.

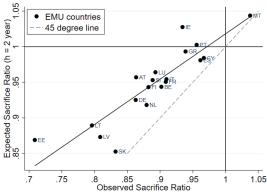
^{***}p < 0.01, **p < 0.05, *p < 0.1.

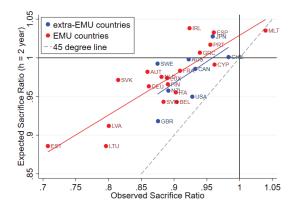
 $[\]P$ Short-term real interest rate forecast errors interacted with a dummy equal to one if the country belongs to the euro area.

Figure 1: Sacrifice ratio across countries



(b) IMF WEO forecasts





Note. The sacrifice ratio is computed here as the ratio between the output loss/gain and the price loss/gain over a two-year horizon (100=2021). Fixed event data. The gray dashed line represents the 45-degree line. The vertical axis shows the expected sacrifice ratio two years ahead, based on the April 2022 IMF WEO forecast (Fig.1a) and the June ECB/ESCB forecast (Fig.1b). The horizontal axis shows the observed sacrifice ratio for 2024, based on the April 2024 IMF WEO forecast (Fig.1a) and the June 2024 ECB/ESCB forecast (Fig.1b).

Table 3 focuses on IMF WEO forecasts, which cover a larger group of countries compared to the ECB/Eurosystem: 29 advanced economies, 19 of which belong to the euro area. We interact the monetary policy stance variable with a dummy indicator equal to one for euro area countries. While β_R is not significant in all specifications, the interaction is always significant, indicating that for the IMF forecast, the bias in monetary multipliers comes only from the euro area. In other words, IMF forecasters correctly incorporated the impact of monetary policy for extra-euro area countries but missed the monetary multiplier for euro area countries, similarly to ECB/Eurosystem staff.

In summary, our findings for the euro area indicate that expected monetary policy tightening have been associated with (i) higher GDP growth and (ii) lower inflation than projected. This implies that forecasters overestimated the sacrifice ratio – the cost of output lost per 1\% reduction in inflation – during the last monetary policy tightening. Figure 1 illustrates this using IMF WEO and ECB/Eurosystem projections. The vertical axis shows the projected sacrifice ratio before the monetary tightening began in July 2022 (IMF: April 2022, ECB: June 2022), while the horizontal axis shows the observed sacrifice ratio based on the latest data (IMF: April 2024, ECB: June 2024). Comparing the linear fits for EMU and extra-EMU countries with the 45-degree line (grey dashed line), we see that most points lie above this line. This indicates that the observed sacrifice ratio was lower than expected, suggesting that forecasters overestimated the costs of disinflation.

5 Robustness analysis

We carry out several robustness analyses, reported in Section C of the Supplementary Appendix. First, we consider the one-year horizon or pool one-year and two-year horizons to increase sample size, which is limited in our setting (see Tables C.1 and C.2 for GDP and inflation, respectively).

Second, we switch to the fixed-event setting, performing the analysis (i) controlling for quarter of origin dummies, or (ii) considering only one specific quarter of origin at a time. Both options ensure comparability across forecast exercises, although option (ii) drastically reduces observations. Results based on ECB/Eurosystem forecasts are in Tables C.3 and C.4 for GDP and inflation, respectively. Tables C.5 and C.6 display the same results based on IMF WEO projections.

Third, we consider a subset of ECB projection exercises where all observed values of variables are available, avoiding the need to proxy them with the last available forecast of June 2024 (see Tables C.7 and C.8).

Fourth, we consider alternative proxies for monetary policy stance: (i) short-term nominal interest rates as in Forbes et al. (2024) (see Tables C.9 and C.10); (ii) add in the specification long-term real interest rates, which provide insights into agents' perceptions of the future path of real short-term rates (see Table C.11).

Fifth, we add country fixed effects to the baseline analysis (see Tables C.12 and C.13 for GDP growth and inflation, respectively). This corresponds to estimating different average forecast errors for each country, implying that forecast errors can vary across countries.

Finally, we carry out a robustness check which includes euro area GDP growth forecast error as additional control variable. This would help account for the potential endogeneity between country-specific GDP growth forecast errors and the monetary policy stance. Results are robust and are reported in Table C.14 of the Appendix.

All results confirm that our main findings are robust to alternative specifications, sample selections, and definitions of monetary policy stance. The only exception is the analysis based on the fixed-event setting, where results for GDP confirm our baseline results, but results for inflation are not always consistent and sometimes change sign. Given the comparability limitations and small sample size of fixed events, we take the results from this robustness exercise with caution.

6 Conclusions

This paper studies the role of monetary policy during this last tightening cycle in the euro area. We find a positive association between monetary policy stance forecasts and growth forecast errors. This suggests that the effect of monetary policy has been milder than forecasters assumed, implying it may have been less effective than initially thought. In inflation regressions instead, the coefficient associated with monetary policy stance is negative and significant, suggesting that that forecasters have underestimated monetary multipliers. These results are robust across various specifications, robustness checks, and forecasting exercises (ECB/Eurosystem and IMF forecasts). Finally, we find that IMF forecasters have incorrectly quantified the monetary multiplier only for euro area countries, not for extra-EMU countries. Putting all this together, our interpretation is that for euro area countries, the sacrifice ratio, or disinflation cost, has been lower than expected.

We rationalize our findings based on several factors. Forecasters may have faced uncertainty regarding new elements that reduced the impact of monetary policy on GDP, such as excess savings from the COVID-19 period, ample liquidity in the financial system, and households' fixed-term rate indebtedness accompanied by private deleveraging in the euro area. Additionally, the underestimation of the sacrifice ratio aligns with the hypothesis that the implied Phillips curve is steeper than assumed in forecasts. Estimating the slope of the Phillips curve is uncertain. Recent policy discussions have questioned the validity of the latest Phillips curve estimations, which suggested a relatively flat curve and a high sacrifice ratio. However, factors such as anchored inflation expectations and accelerated repricing by firms, observed during the recent monetary tightening cycle, align with a steeper Phillips curve (Escriva, 2024).

Our findings suggest that forecasters may have inaccurately perceived the recent inflationary crisis in the euro area as predominantly supply-driven, underestimating its demand-driven component. This led to the belief that monetary policy in the euro area would be exceedingly costly in terms of output. However, recent evidence, including that presented by Giannone and Primiceri (2024), suggests that demand factors also significantly contribute to inflation dynamics.

References

- Ball, Laurence. (1994). "What Determines the Sacrifice Ratio?". NBER Working Paper Series, 4306, National Bureau of Economic Research. https://doi.org/10.3386/w4306
- Benigno, Gianluca, Boris Hofmann, Galo Nuño Barrau and Damiano Sandri. (2024). "Quo vadis, r*? The natural rate of interest after the pandemic". *BIS Quarterly Review*, March, Bank for International Settlements. https://ideas.repec.org/a/bis/bisqtr/2403b.html
- Blanchard, Olivier J., and Daniel Leigh. (2013). "Growth forecast errors and fiscal multipliers". *American Economic Review*, 103(3), pp. 117-120. https://doi.org/10.1257/aer.103.3.117
- Blanchard, Olivier J., and Daniel Leigh. (2014). "Learning about fiscal multipliers from growth forecast errors". *IMF Economic Review*, 62(2), pp. 179–212. https://doi.org/10.1057/imfer.2014.17
- Brinca, Pedro, Miguel H. Ferreira, Francesco Franco, Hans A. Holter and Laurence Malafry. (2021). "Fiscal consolidation programs and income inequality". *International Economic Review*, 62(1), pp. 405-460. https://doi.org/10.1111/jere.12482
- Buncic, Daniel. (2024). "Econometric issues in the estimation of the natural rate of interest". *Economic Modelling*, 132(106641). https://doi.org/10.1016/j.econmod.2023.106641
- Chahad, Mohammed, Anna Hofmann-Drahonsky, Catalina Martínez Hernández and Adrian Page. (2024). "An update on the accuracy of recent Eurosystem/ECB staff projections for short-term inflation". ECB Economic Bulletin, 2/2024, European Central Bank. https://www.ecb.europa.eu/press/economic-bulletin/focus/2024/html/ecb.ebbox202402_05~10d8d08f79.en.html
- Darvas, Zsolt. (2018). "Forecast errors and monetary policy normalisation in the euro area". Bruegel Policy Contributions, 28816, Bruegel. https://ideas.repec.org/p/bre/polcon/28816.html
- Dovern, Jonas, Ulrich Fritsche and Jiri Slacalek. (2012). "Disagreement Among Forecasters in G7 Countries". *The Review of Economics and Statistics*, 94(4), pp. 1081-1096. https://doi.org/10.1162/REST_a_00207
- Escrivá, José Luis. (2024). "The makings of an immaculate disinflation" [Summary]. Hydra Conference. https://www.bde.es/f/webbe/GAP/Secciones/SalaPrensa/IntervencionesPublicas/Gobernador/Arc/Fic/IIPP-2024-09-27-escriva-en-or.pdf
- Fatás, Antonio, and Lawrence H. Summers. (2018). "The permanent effects of fiscal consolidations". *Journal of International Economics*, 112, pp. 238-250. https://doi.org/10.1016/j.jinteco.2017.11.007
- Fiorentini, Gabriele, Alessandro Galesi, Gabriel Pérez-Quirós and Enrique Sentana. (2018). "The rise and fall of the natural interest rate". Documentos de Trabajo, 1822, Banco de España. https://doi.org/10.2139/ssrn.3214487
- Forbes, Kristin, Jongrim Ha and M. Ayhan Kose. (2024). "Rate cycles". CEPR Discussion Paper, 19272, Center for Economic Policy Research Press. https://doi.org/10.2139/ssrn.5063421
- Giannone, Domenico, and Giorgio Primiceri. (2024). "The drivers of post-pandemic inflation". NBER Working Paper Series, 32859, National Bureau of Economic Research. https://doi.org/10.3386/w32859

- Granziera, Eleonora, Pirkka Jalasjoki and Maritta Paloviita. (2024). "The bias of the ECB inflation projections: A state-dependent analysis". *Journal of Forecasting*, November. https://doi.org/10.1002/for.3236
- Lane, Philip R. (2024). "Inflation in the euro area" [Speech]. Institute of International and European Affairs. https://www.bis.org/review/r240527c.htm
- Lewis, Christine, and Nigel Pain. (2014). "Lessons from OECD forecasts during and after the financial crisis". *OECD Journal: Economic Studies*, 2014(1), pp. 9-39. https://doi.org/10.1787/eco_studies-2014-5jxrcm2glc7j
- Lubik, Thomas A., Brennan Merone and Nathan Robino. (2024). "Stargazing: Estimating r* in Other Countries". *Richmond Fed Economic Brief*, 24(10). https://ideas.repec.org/a/fip/fedreb/97953. html
- Suyker, Wim, and Jan Möhlmann. (2015). "An update of Blanchard's and Leigh's estimates in 'Growth Forecast Errors and Fiscal Multipliers". CPB Background Document, November, CPB Netherlands Bureau for Economic Policy Analysis. https://www.cpb.nl/sites/default/files/publicaties/download/cpb-background-document-update-blanchard-and-leigh-estimates-growth-forecast-errors-and-fiscal-multi.pdf
- Taylor, John B. (1993). "Discretion versus policy rules in practice". *Carnegie-Rochester Conference Series on Public Policy*, 39, pp. 195-214. https://doi.org/10.1016/0167-2231(93)90009-L

Appendix

A Details on interest rates used in the WEO database

To construct a projection for interest rates, we use interest rate futures, in line with the literature.¹⁵ This section report in detail which interest rate futures we consider for each country.

• Euro area

- Interest rate futures are those of the 3-month EURIBOR, as quoted on EUREX.
- At the beginning of a year, there are interest rate futures for that same year with monthly expiration.¹⁶ For subsequent years, interest rate futures are only available with quarterly expiration, that is, March, June, September, and December. Therefore we take 12 months averages for the current year and 4 months-averages for the following years.

• Japan

- Interest rate futures refer to the 3-month Euroyen rates and TONA ¹⁷ as listed on 'Tokyo-Financial'.
- Interest rate futures consistently expire on a quarterly basis. For this reason, we take 4 months-averages for each year.

• United Kingdom

- Interest rate futures are based on the 3-month SONIA as quoted on Euronext Liffe.
- Interest rate futures typically expire on a quarterly basis. As a result, in the majority of instances, the averaging process is conducted every quarter.

• United States

 Interest rate futures are based on the 3-month SOFR as quoted on the Chicago Mercantile Exchange.

 $^{^{15}}$ Interest rate futures are financial contracts between the buyer and seller agreeing to the future delivery of an interest-bearing asset.

¹⁶Except for those expiring in July (trading begins in mid-January), August (trading begins in mid-February), September (trading starts in April), and October (trading starts in May).

¹⁷For the 2023 WEO, the rates are those of the 3-month TONA, which are quoted in the same source as the previous ones.

The interest rate futures have quarterly expiration until 2022, but from 2023 onwards, they are monthly. The method of averaging is largely akin to the approach used for the United Kingdom.

• Australia

- The closest futures interest rates are those based on 90-day bank bills, as quoted on ASX Trade24.
- Interest rate futures are set to expire every quarter.

• Canada

- The future interest rate refers to the interest rate of 90-day bank deposits, as quoted in Montreal.
- Until 2022 (included), the expiration of interest rate futures was set on a monthly basis. However, a shift to a quarterly expiration cycle was implemented starting from the end of 2023.

• New Zealand

- The interest rate futures are based on the 90-day bank bill rate as quoted on ASX
 Trade24.
- The interest rate futures have quarterly expiration, so it is averaged every three months and rates are only available for the current year and the following two years.

• Switzerland

- The interest rate futures is based on the 3-month SARON as quoted on ICE.
- Interest rate futures contracts have quarterly expirations, and data is available
 for the current year and the following three years.

• Sweden

- Interest rate futures are on the 3-month STIBOR, as quoted on Nasdaq.
- Futures contracts expire quarterly, and there is data available for the current year
 and the following two years.

B Fixed-Event and Fixed-Horizon Forecasts

Suppose that one wants to compare y + 1 GDP forecast realized in the first quarter of year y with those published in the subsequent quarters of year y. The comparison is not fully correct because the information set within the year used to elaborate the forecast differs. More formally, since the quarterly forecast exercises target yearly variables (GDP growth and inflation), the different forecasts that are elaborated along the same calendar year are not comparable since the information set changes over time, with uncertainty decreasing over the calendar year. This setting is known as fixed-event forecast.¹⁸

As a first approximation, one can tackle this problem by controlling for the information set used to formulate the forecast. In our case, this means controlling for quarter dummies of forecast origin. Alternatively, we can estimate the model by the quarter in which each projection exercise is conducted. We include both exercises in our list of robustness checks.

A more correct way to tackle this issue is to move away from the fixed-event forecast setting towards the fixed-horizon forecasts, in which one compares forecasts with the same forecast horizon and thus, comparable information sets. The difference between the two settings is shown in Figure B.1, in an example for two projection exercises (carried out in the first and second quarter of year y) at one-year horizon (y + 1). In both panels, the arrows represents the horizon of each projection exercise, and its corresponding area reflects the actual horizon implied from the exercise. In the fixed-event setting (Figure B.1a), the forecast exercise carried out in the second quarter implies a one quarter shorter horizon than the exercise closed one quarter before. This boils down to comparability issues. By contrast, the fixed-horizon setting (Figure B.1b) ensures horizons with the same length across exercises, which restores the comparability across forecasts.

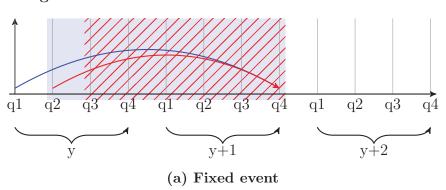
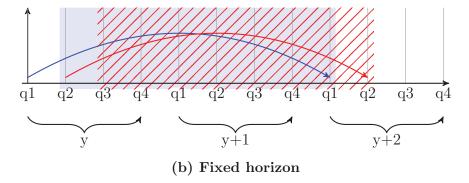


Figure B.1: Fixed-events vs fixed horizon forecasts

¹⁸In a fixed-event forecast, the researcher fixes the target event to forecast - in this case y + 1 GDP growth - and compares projections realized at different points in time.



Note: The arrow represents the projection exercise. The begin of the arrow indicates the origin of the forecast, and the end of the arrow indicates the forecast target (horizon). The figure only shows, as an example, two projection exercises at one-year horizon. The blue (red) arrow represents the forecast exercise carried out in q1 (q2) of year y to forecast the outlook one-year ahead (y + 1). The light blue (red) area represents the horizon ahead in q1 (q2) of year y.

To move to fixed horizon setting, we follow Dovern et al. (2012), and approximate the fixed horizon forecast for the next 8 quarters (two years, i.e. the medium term horizon) as an average of the forecasts for the current and next two calendar years weighted by their share in forecasting horizon, according to the following equation, in which q correspond with the projection exercise:

$$\tilde{\hat{x}}_{y+2|y} = \frac{4-q}{8}\hat{x}_{y|y} + \frac{4}{8}\hat{x}_{y+1|y} + \frac{q}{8}\hat{x}_{y+2|y}$$
(B.1)

In particular, say that, in year y, we dispose of (fixed-event) 4 different quarterly GDP forecasts for y + 1 and y + 2, elaborated in each quarter q of y. The fixed-horizon forecast is fixed at two years (8 quarters), and is represented by the shaded areas in Figure B.1b. Hence, the fixed-horizon forecast for the projection exercise carried out in the first quarter of y is given by the weighted average of the fixed-event forecasts for the current year y ($\hat{x}_{y|y}$) and for the two subsequent years ($\hat{x}_{y+1|y}$ and $\hat{x}_{y+2|y}$), weighted by the corresponding share in the forecast horizon.

$$\begin{split} \tilde{\hat{x}}_{y+2|y,q1} &= \frac{3}{8}\hat{x}_{y|y} + \frac{4}{8}\hat{x}_{y+1|y} + \frac{1}{8}\hat{x}_{y+2|y} \\ &[\ldots] \\ \tilde{\hat{x}}_{y+2|y,q4} &= \frac{1}{8}\hat{x}_{y|y} + \frac{4}{8}\hat{x}_{y+1|y} + \frac{3}{8}\hat{x}_{y+2|y} \end{split}$$

Note that the fixed-horizon setting ensures that different projections of the same calendar year share the same horizon. This improves comparability across projections. Nevertheless,

the information set upon which projections are formulated changes along the calendar year.¹⁹ To account for this possible source of bias, we additionally control for quarter of origin dummies in the fixed-horizon setting.

¹⁹In addition, there could be specific features associated with projection exercises of specific quarters. For instance, first and third quarter forecasts are carried out by the ECB, as opposed to the other projections, which are jointly realized by all Eurosystem central banks.

C Robustness

Table C.1: Annual GDP growth forecast errors on monetary policy stance on at different horizons

				ECB/ESC	B Forecas	ts						IMF WE	O Forecas	ts		
		1 year	ahead			1 & 2 ye	ears ahead			1 year	ahead			1 & 2 yea	ars ahead	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
Real R^{\dagger}	0.142*** (0.001)	0.149*** (0.001)	0.163*** (0.001)	0.167*** (0.003)	0.173***	0.175*** (0.000)	0.186*** (0.000)	0.181*** (0.000)	0.192** (0.018)	0.226*** (0.008)	0.222*** (0.004)	0.338**	0.231***	0.272*** (0.000)	0.262*** (0.000)	0.312*** (0.001)
$\mathrm{Fiscal}^{\ddagger}$	(0.001)	0.099	0.107	0.125	(0.000)	0.027	0.032	0.040 (0.689)	(0.010)	0.448*	0.452*	0.433*	(0.000)	0.462**	0.475**	0.463**
Energy [§]		(0.002)	-0.004 (0.467)	-0.007 (0.359)		(0.100)	-0.003 (0.440)	-0.003 (0.636)		(0.001)	0.000 (0.907)	-0.005 (0.361)		(0.012)	0.001 (0.574)	-0.001 (0.808)
q=2			(0.101)	0.003			(0.110)	0.023			(0.001)	(0.001)			(0.011)	(0.000)
q=3				-0.282 (0.461)				-0.005 (0.981)								
q = 4				0.066				0.172 (0.266)								
half=2				(0.020)				(0.200)				-0.803 (0.122)				-0.325 (0.279)
Constant	-0.121 (0.391)	-0.145 (0.291)	-0.175 (0.209)	-0.164 (0.405)	-0.231** (0.022)	-0.237** (0.012)	-0.262*** (0.006)	-0.310*** (0.010)	-0.204 (0.363)	-0.328 (0.159)	-0.308 (0.253)	-0.005 (0.989)	-0.302** (0.024)	-0.425*** (0.003)	-0.364** (0.021)	-0.251 (0.244)
servations R-squared	171 0.0757	171 0.0803	171 0.0844	171 0.0904	342 0.0988	342 0.0990	342 0.101	342 0.104	76 0.0654	76 0.142	76 0.142	76 0.168	152 0.0892	152 0.162	152 0.164	152 0.170

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between GDP growth observed and forecasted). Two-years ahead horizon considered. The sample refers to euro area countries (19).

† Real R refers to real short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

\$ Energy refers to energy price forecast errors.

* * *p < 0.01, ** p < 0.05, *p < 0.1.

Table C.2: Annual inflation forecast errors on monetary policy stance at different horizons

				ECB/ESC	B Forecasts							IMF WE	Forecasts			
	(1)	1 year (2)	ahead (3)	(4)	(5)	1 & 2 ye (6)	ars ahead (7)	(8)	(9)	1 year (10)	ahead (11)	(12)	(13)	1 & 2 yes (14)	ars ahead (15)	(16)
Real R^{\dagger}	-0.286***	-0.264*** (0.000)	-0.394*** (0.000)	-0.381***	-0.291***	-0.268*** (0.000)	-0.397*** (0.000)	-0.369***	-0.499***	-0.462*** (0.000)	-0.493*** (0.000)	-0.144 (0.366)	-0.482*** (0.000)	-0.438*** (0.000)	-0.450*** (0.000)	-0.141 (0.235)
$Fiscal^{\ddagger}$	(0.000)	0.291*	0.212	0.195 (0.172)	(0.000)	0.316**	0.250*	0.223*	(0.000)	0.486** (0.022)	0.522**	0.465**	(0.000)	0.490*** (0.003)	0.505***	0.434*** (0.002)
Energy§		(0.000)	0.040***	0.032***		(0.020)	0.036***	0.026***		(0.022)	0.004 (0.398)	-0.012** (0.023)		(0.000)	0.002	-0.011*** (0.001)
q = 2			()	-0.882*** (0.000)			(/	-0.754*** (0.000)			(,	(/			(/	()
q = 3				-0.903*** (0.004)				-0.831*** (0.000)								
q = 4				-0.865*** (0.002)				-0.971*** (0.000)								
half=2												-2.418*** (0.000)				-1.995*** (0.000)
Constant	-0.182* (0.066)	-0.251** (0.026)	0.031 (0.777)	0.538*** (0.001)	-0.110 (0.104)	-0.191** (0.015)	0.099 (0.244)	0.593*** (0.000)	-0.382** (0.036)	-0.517*** (0.007)	-0.360 (0.178)	0.552* (0.069)	-0.175 (0.104)	-0.306*** (0.005)	-0.234 (0.164)	0.460** (0.012)
Observations R-squared	171 0.204	171 0.230	171 0.472	171 0.516	342 0.188	342 0.213	342 0.414	342 0.462	76 0.254	76 0.306	76 0.315	76 0.450	152 0.240	152 0.290	152 0.292	152 0.423

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between inflation observed and forecasted). Two-years ahead horizon considered. The sample refers to euro area countries (19).

† Real R refers to real short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

\$ Energy refers to energy price forecast errors.

* * *p < 0.01, * * p < 0.05, *p < 0.1.

			2 у	ears ahead						1 & :	years ahea	nd		
	(1) Full sample	(2) Full sample	(3) Full sample	(4) Origin Q1	(5) Origin Q2	(6) Origin Q3	(7) Origin Q4	(8) Full sample	(9) Full sample	(10) Full sample	(11) Origin Q1	(12) Origin Q2	(13) Origin Q3	(14) Origin Q4
Real R^{\dagger}	0.434*** (0.000)	0.419*** (0.000)	0.249** (0.021)	0.448*** (0.002)	0.667***	0.645*** (0.000)	0.222 (0.122)	0.254*** (0.002)	0.233*** (0.004)	0.180** (0.045)	0.705*** (0.000)	0.226* (0.094)	0.136 (0.143)	0.119 (0.189)
Fiscal [‡]	(01000)	-0.085 (0.524)	-0.115 (0.381)	0.196 (0.277)	-0.168 (0.524)	-0.064 (0.741)	-0.303 (0.135)	(0.00-)	-0.257** (0.019)	-0.259** (0.018)	-0.348* (0.077)	-0.427** (0.021)	-0.157 (0.284)	-0.077 (0.615)
Energy§			0.023*** (0.000)							0.008 (0.222)				
q = 2	(0.001)	-0.415*** (0.002)	-0.267** (0.013)					-0.169 (0.210)	-0.159 (0.239)	-0.126 (0.336)				
q = 3	-0.462*** (0.002)	-0.459*** (0.002)	0.269* (0.067)					0.198 (0.230)	0.215 (0.198)	0.461** (0.021)				
q = 4	-0.260 (0.181)	-0.228 (0.238)	0.082 (0.614)					0.492*** (0.007)	0.485*** (0.007)	0.572*** (0.000)				
Constant	-0.290*** (0.001)	-0.287*** (0.001)	-0.270*** (0.003)	-0.298*** (0.000)	-0.708*** (0.000)	-0.809*** (0.000)	-0.438*** (0.005)	-0.688*** (0.000)	-0.628*** (0.000)	-0.617*** (0.000)	-0.468*** (0.000)	-0.738*** (0.000)	-0.522*** (0.000)	-0.266** (0.021)
Observations R-squared	171 0.189	171 0.192	171 0.281	57 0.252	38 0.227	38 0.247	38 0.122	342 0.118	342 0.138	342 0.143	114 0.344	76 0.160	76 0.067	76 0.037

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between GDP growth observed and forecasted). Fixed-event structure considered. The sample refers to euro area countries (19).

† Real R refers to real short term interest rate forecast.

‡ Fiscal refers to energy price forecast errors.

* * *p < 0.01, * * p < 0.05, *p < 0.1.

-			2 y	ears ahead						1 & :	2 years ahea	ad		
	(1) Full sample	(2) Full sample	(3) Full sample	(4) Origin Q1	(5) Origin Q2	(6) Origin Q3	(7) Origin Q4	(8) Full sample	(9) Full sample	(10) Full sample	(11) Origin Q1	(12) Origin Q2	(13) Origin Q3	(14) Origin Q4
Real R [†]	0.269** (0.019)	0.234* (0.083)	0.263* (0.080)	-0.191* (0.077)	-0.005 (0.974)	0.167 (0.205)	0.817*** (0.000)	-0.112* (0.088)	-0.105 (0.102)	-0.065 (0.442)	-0.828*** (0.000)	-0.266 (0.119)	0.020 (0.797)	0.188*** (0.000)
Fiscal [‡]		-0.214 (0.128)	-0.208 (0.134)	-0.328** (0.043)	-0.324 (0.189)	-0.262 (0.131)	0.014 (0.903)		0.087 (0.461)	0.088 (0.451)	-0.002 (0.994)	0.268 (0.270)	0.059 (0.755)	-0.191** (0.041)
Energy§		()	-0.004 (0.490)	()	()	(/	()		()	-0.006 (0.341)	()	(* **)	()	()
q = 2	-0.042 (0.691)	-0.043 (0.679)	-0.069 (0.483)					-0.249 (0.117)	-0.252 (0.114)	-0.277* (0.078)				
$\mathbf{q}=3$	-0.226* (0.084)	-0.220* (0.097)	-0.344* (0.057)					-0.901*** (0.000)	-0.906*** (0.000)	-1.092*** (0.000)				
q = 4	-0.694*** (0.000)	-0.613*** (0.002)	-0.666*** (0.001)					-1.142*** (0.000)	-1.139*** (0.000)	-1.205*** (0.000)				
Constant	0.224** (0.018)	0.231** (0.014)	0.228** (0.018)	0.198** (0.012)	0.196 (0.122)	0.030 (0.821)	-0.432*** (0.001)	0.581*** (0.000)	0.561*** (0.000)	0.552*** (0.000)	0.374*** (0.000)	0.161 (0.247)	-0.243* (0.054)	-0.311*** (0.001)
Observations R-squared	171 0.186	171 0.208	171 0.211	57 0.095	38 0.057	38 0.090	38 0.661	342 0.118	342 0.120	342 0.123	114 0.287	76 0.140	76 0.002	76 0.194

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between inflation observed and forecasted). Fixed-event structure considered. The sample refers to euro area countries (19).

† Real R refers to real short term interest rate forecast.

‡ Fixed refers to fiscal stance forecast.

^{*} Energy refers to energy price forecast errors. * * * p < 0.01, * * p < 0.05, * p < 0.1. † Energy price forecast error.

Table C.5: Annual GDP growth forecast errors on monetary policy stance at different horizons. Fixed event data. IMF WEO forecasts.

		2	years ahead				1 &	2 years ahea	ıd	
	(1) Full sample	(2) Full sample	(3) Full sample	(4) Origin H1	(5) Origin H2	(6) Full sample	(7) Full sample	(8) Full sample	(9) Origin H1	(10) Origin H2
Real R^{\dagger}	0.555*** (0.000)	0.601*** (0.000)	0.453*** (0.005)	0.622** (0.012)	0.546*** (0.001)	0.347*** (0.003)	0.396*** (0.007)	0.308** (0.032)	0.958*** (0.000)	0.236 (0.130)
Fiscal [‡]	. ,	(0.603)	0.257 (0.513)	(0.820)	(0.508)	, ,	0.300 (0.212)	(0.317	0.528* (0.061)	0.288 (0.329)
Energy§			0.009*** (0.001)					0.005* (0.089)		
half=2	-0.264 (0.109)	-0.272 (0.100)	0.246 (0.128)			0.399** (0.027)	0.446** (0.014)	0.705*** (0.002)		
Constant	-0.625*** (0.000)	-0.669*** (0.000)	-0.448*** (0.003)	-0.639*** (0.000)	-0.925*** (0.000)	-0.867*** (0.000)	-0.988*** (0.000)	-0.862*** (0.000)	-0.827*** (0.000)	-0.581*** (0.001)
Observations R-squared	76 0.140	76 0.147	76 0.274	38 0.181	38 0.106	152 0.135	152 0.158	152 0.175	76 0.312	76 0.092

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between GDP growth observed and forecasted). Fixed-event structure considered. The sample refers to euro area countries (19).

† Read R refers to real short term interest rate forecast.

 $^{^{\}ddagger}$ Fiscal refers to fiscal stance for ecast.

[§] Energy refers to energy price forecast errors. ****p < 0.01, **p < 0.05, *p < 0.1.

Table C.6: Annual inflation forecast errors on monetary policy stance at different horizons. Fixed event data. IMF WEO forecasts.

		2	years ahead				1 &	2 years ahea	ıd	
	(1) Full sample	(2) Full sample	(3) Full sample	(4) Origin H1	(5) Origin H2	(6) Full sample	(7) Full sample	(8) Full sample	(9) Origin H1	(10) Origin H2
Real R^{\dagger}	0.515***	0.514***	0.608***	0.366	0.729***	-0.118	-0.055	0.205	-0.520	0.078
	(0.000)	(0.001)	(0.000)	(0.109)	(0.000)	(0.302)	(0.646)	(0.146)	(0.109)	(0.470)
Fiscal [‡]		-0.003	-0.046	-0.013	-0.041		0.381**	0.329*	0.339	0.236
Energy§		(0.992)	(0.859) -0.006** (0.027)	(0.979)	(0.831)		(0.042)	(0.068) -0.015*** (0.000)	(0.246)	(0.192)
half=2	-0.603*** (0.001)	-0.603*** (0.001)	-0.931*** (0.000)			-0.900*** (0.000)	-0.840*** (0.000)	-1.604*** (0.000)		
Constant	0.249* (0.098)	0.249 (0.119)	0.110 (0.465)	0.231 (0.200)	-0.474*** (0.004)	0.646*** (0.001)	0.493*** (0.007)	0.123 (0.476)	0.290* (0.088)	-0.269** (0.035)
	(0.098)	(0.119)	(0.400)	(0.200)	(0.004)	(0.001)	(0.007)	(0.470)	(0.000)	(0.033)
Observations	76	76	76	38	38	152	152	152	76	76
R-squared	0.160	0.160	0.222	0.077	0.286	0.0930	0.125	0.245	0.128	0.034

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between inflation observed and forecasted). Fixed-event structure considered. The sample refers to euro area countries (19).

† Real R refers to real short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

[§] Energy refers to energy price forecast errors. ***p < 0.01, **p < 0.05, *p < 0.1.

Table C.7: Annual GDP growth errors on monetary policy stance only with observed values.

		ECB/ESC	B Forecats	5	I	MF WEO		s
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real R [†]	0.346*** (0.001)	0.362*** (0.001)	0.362*** (0.002)	0.390*** (0.002)	0.382*** (0.009)	0.381** (0.013)	0.446**	0.446**
$Fiscal^{\ddagger}$	(0.001)	0.113	0.108	0.091	(0.000)	0.349	0.322	0.322
Energy [§]		(0.465)	(0.489) 0.002 (0.781)	(0.582) 0.005 (0.714)		(0.330)	(0.347) 0.008 (0.234)	(0.347) 0.008 (0.234)
q = 2			(0.781)	0.418			(0.234)	(0.234)
q = 3				(0.172) 0.470 (0.601)				
q = 4				-				
half=2								-
Constant	1.053** (0.033)	1.061** (0.031)	1.106** (0.024)	1.117** (0.038)	0.584 (0.310)	0.332 (0.626)	1.374 (0.187)	1.374 (0.187)
Observations R-squared	76 0.190	76 0.196	76 0.197	76 0.211	38 0.141	38 0.181	38 0.213	38 0.213

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between GDP growth observed and forecasted). Fixed-horizon structure considered. The sample refers to euro area countries (19). Based on subset of projection exercises in which all observed values of variables are available and there is not need to proxy them with the last available forecast of June 2024.

† Real R refers to real short term interest rate forecast.

‡ Fixed refers to fiscal stance forecast.

[§] Energy refers to energy price forecast errors. ***p < 0.01, **p < 0.05, *p < 0.1.

Table C.8: Annual inflation forecast errors on monetary policy stance only with observed values.

	1		CB Forecat	ts			O Forecasts s ahead	s
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real R [†]	-0.333** (0.035)	-0.258 (0.133)	-0.237* (0.051)	-0.154 (0.183)	-0.428** (0.014)	-0.430** (0.010)	-0.133 (0.355)	-0.133 (0.355)
$Fiscal^{\ddagger}$	()	0.542* (0.055)	0.369**	0.154 (0.374)	()	0.514 (0.113)	0.391*	0.391* (0.054)
Energy [§]		()	0.050*** (0.000)	0.094*** (0.000)		()	0.036*** (0.000)	0.036*** (0.000)
q = 2			(,	-0.711** (0.033)			()	()
q = 3				3.184*** (0.000)				
q = 4				- ′				
half=2								-
Constant	-0.197 (0.759)	-0.160 (0.808)	1.329*** (0.003)	2.350*** (0.000)	0.381 (0.476)	0.011 (0.984)	4.762*** (0.000)	4.762*** (0.000)
Observations R-squared	76 0.0883	76 0.153	76 0.585	76 0.698	38 0.117	38 0.174	38 0.613	38 0.613

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between inflation observed and forecasted). Fixed-horizon structure considered. The sample refers to euro area countries (19). Based on subset of projection exercises in which all observed values of variables are available and there is not need to proxy them with the last available forecast of June 2024.

† Real R refers to real short term interest rate forecast.

 $^{^{\}ddagger}$ Fiscal refers to fiscal stance for ecast.

[§] Energy refers to energy price forecast errors. ****p < 0.01, **p < 0.05, *p < 0.1.

Table C.9: Annual GDP growth forecast errors on monetary policy stance using nominal short term interest rate.

			s ahead			IMF WEO	ahead	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Nominal ST rate [†]	0.289*** (0.000)	0.278*** (0.000)	0.306***	0.306***	0.269* (0.053)	0.364** (0.015)	0.356** (0.015)	0.253 (0.206)
Fiscal [‡]	()	-0.142 (0.409)	-0.132 (0.443)	-0.127 (0.468)	()	0.444 (0.156)	0.498 (0.112)	0.496 (0.115)
Energy§			-0.006 (0.286)	-0.006 (0.405)			0.005 (0.176)	0.006 (0.115)
q = 2				0.069 (0.673)				
q = 3				-0.003 (0.989)				
q = 4				0.080 (0.727)				
half=2								0.283 (0.352)
Constant	-0.265** (0.024)	-0.226** (0.030)	-0.273** (0.010)	-0.309** (0.035)	-1.391*** (0.003)	-1.803*** (0.001)	-1.542*** (0.002)	-1.325** (0.017)
Observations R-squared	171 0.198	171 0.203	171 0.210	171 0.211	76 0.0427	76 0.0955	76 0.128	76 0.131

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between GDP growth observed and forecasted). The sample refers to euro area countries (19). Monetary policy stance approximated with nominal short term interest rate.

† Nominal ST R refers to nominal short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

‡ Energy refers to energy price forecast errors.

* * * p < 0.01, * p < 0.05, * p < 0.1.

Table C.10: Annual inflation forecast errors on monetary policy stance using nominal short term interest rate.

	ECB/ESCB Forecats 2 years ahead				IMF WEO Forecasts 2 years ahead				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	
Nominal ST R [†]	-0.324*** (0.000)	-0.297*** (0.001)	-0.459*** (0.000)	-0.418*** (0.000)	-1.306*** (0.000)	-1.262*** (0.000)	-1.258*** (0.000)	-1.502*** (0.000)	
$Fiscal^{\ddagger}$	(0.000)	0.373	0.316	0.277 (0.201)	(0.000)	0.202	0.171 (0.244)	0.166	
Energy [§]		(0.122)	0.033***	0.025*** (0.001)		(0.110)	-0.003 (0.200)	0.001 (0.699)	
$\mathbf{q}=2$			(0.000)	-0.652*** (0.001)			(0.200)	(0.000)	
q = 3				-0.624** (0.011)					
q = 4				-0.957*** (0.001)					
half=2								0.671*** (0.005)	
Constant	-0.069	-0.170*	0.104	0.553***	3.882***	3.695***	3.543***	4.056***	
	(0.390)	(0.094)	(0.359)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	
Observations	171	171	171	171	76	76	76	76	
R-squared	0.172	0.196	0.370	0.422	0.689	0.697	0.704	0.718	

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between inflation observed and forecasted). The sample refers to euro area countries (19). Monetary policy stance approximated with nominal short term interest rate.

† Nominal ST R refers to nominal short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

‡ Energy refers to energy price forecast errors.

* * *p < 0.01, *p < 0.05, *p < 0.1.

Table C.11: Annual GDP growth and inflation forecast errors on monetary policy stance on using real short and long term interest rate.

					CB Foreca	ts			
	GDP growth			inflation					
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	
Real ST R [†]	0.240***	0.221**	0.248**	0.252**	-0.030	0.032	-0.129	-0.133	
	(0.003)	(0.013)	(0.012)	(0.019)	(0.703)	(0.679)	(0.178)	(0.194)	
Real LT R¶	0.167	0.188	0.170	0.173	-1.004***	-1.073***	-0.967***	-0.916***	
	(0.156)	(0.144)	(0.208)	(0.230)	(0.000)	(0.000)	(0.000)	(0.000)	
Fiscal [‡]		-0.179	-0.167	-0.167		0.580***	0.514***	0.486***	
		(0.327)	(0.366)	(0.382)		(0.001)	(0.001)	(0.002)	
Energy§			-0.004	-0.006			0.026***	0.024***	
			(0.416)	(0.406)			(0.000)	(0.000)	
q = 2				-0.006				-0.257	
				(0.974)				(0.108)	
q = 3				-0.088				-0.175	
				(0.716)				(0.307)	
q = 4				-0.024				-0.407**	
				(0.919)				(0.046)	
Constant	-0.470**	-0.448**	-0.464**	-0.452**	1.166***	1.095***	1.189***	1.309***	
	(0.020)	(0.016)	(0.012)	(0.018)	(0.000)	(0.000)	(0.000)	(0.000)	
Observations	171	171	171	171	171	171	171	171	
R-squared	0.209	0.217	0.222	0.222	0.452	0.509	0.617	0.626	

Note: Robust pval in parentheses. The dependent variable is annual GDP growth and inflation forecast error (difference between observed and forecasted). The sample refers to euro area countries (19). Monetary policy stance approximated with both real short term and long term interest rate.

† Real ST R refers to real short term interest rate forecast errors. ¶ Real LT R refers to real long term interest rate forecast.

 $^{^{\}ddagger}$ Fiscal refers to fiscal stance forecast.

[§] Energy refers to energy price forecast errors. ****p < 0.01, **p < 0.05, *p < 0.1.

Table C.12: Annual GDP growth forecast errors on monetary policy stance on including country fixed effects.

	ECB/ESCB Forecasts 2 years ahead				IMF WEO Forecasts 2 years ahead			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Real ST R [†]	0.269***	0.262***	0.295***	0.290***	0.274***	0.263***	0.241***	0.179*
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.001)	(0.099)
Fiscal [‡]		-0.084	-0.062	-0.049		-0.064	-0.011	-0.000
		(0.687)	(0.764)	(0.809)		(0.722)	(0.956)	(1.000)
Energy§			-0.005*	-0.004			0.002*	0.004
			(0.059)	(0.246)			(0.093)	(0.143)
q = 2				0.071				
				(0.545)				
q = 3				0.041				
				(0.822)				
q = 4				0.122				
				(0.533)				
half=2								0.180
								(0.551)
Observations	171	171	171	171	76	76	76	76
R-squared	0.702	0.703	0.709	0.710	0.845	0.846	0.853	0.854

Note: Robust pval in parentheses. The dependent variable is annual GDP growth forecast error (difference between observed and forecasted). The sample refers to euro area countries (19). Country fixed effects included (not reported in the table).

† Real ST R refers to real short term interest rate forecast.

‡ Fiscal refers to fiscal stance forecast.

‡ Energy refers to energy price forecast errors.

* ***p < 0.01, ***p < 0.05, *p < 0.1.

$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$		ECB/ESCB Forecasts 2 years ahead				IMF WEO Forecasts 2 years ahead			
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$		(1)			(4)	(5)			(8)
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Real ST R [†]							-0.677*** (0.000)	-0.364 (0.236)
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Fiscal [‡]	, ,				, ,		1.089** (0.020)	1.037**
$\begin{array}{c} q=2 \\ q=3 \\ q=3 \\ q=4 \\ \end{array} \begin{array}{c} -0.673^{***} \\ (0.000) \\ -0.251 \\ (0.240) \\ -0.629^{**} \\ (0.017) \end{array}$	Energy§		(,	0.038***	0.037***		()	0.002	-0.004 (0.416)
q = 4 (0.240) -0.629** (0.017)	q = 2			, ,				, ,	` ′
(0.017)	q = 3				(0.240)				
half=2	_								
	half=2								-0.907 (0.163)
Observations 171 171 171 171 76 76 76 R-squared 0.272 0.332 0.551 0.588 0.497 0.552 0.558									76 0.574

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between observed and forecasted). The sample refers to euro area countries (19). Country fixed effects included (not reported in the table).

† $Real\ ST\ R$ refers to real short term interest rate forecast.

‡ $Fiscal\ refers$ to fiscal stance forecast.

[§] Energy refers to energy price forecast errors. ***p < 0.01, **p < 0.05, *p < 0.1.

Table C.14: Annual GDP growth forecast errors on monetary policy stance including euro area GDP growth forecast error.

		2 years	B Forecast s ahead	
	(1)	(2)	(3)	(4)
Real R [†]	0.346***	0.334***	0.337***	0.332***
	(0.000)	(0.000)	(0.000)	(0.000)
$\mathrm{Fiscal}^{\ddagger}$, ,	-0.120	-0.119	-0.121
		(0.490)	(0.493)	(0.488)
Energy^\S			-0.002	0.005
			(0.658)	(0.464)
q = 2				-0.261
				(0.267)
q = 3				0.169
				(0.512)
q = 4				0.074
				(0.750)
euro area GDP growth forecast error	1.002*	0.960	0.841*	1.523*
	(0.088)	(0.108)	(0.078)	(0.054)
Constant	-1.906**	-1.803*	-1.624**	-2.691**
	(0.048)	(0.066)	(0.044)	(0.042)
Observations	171	171	171	171
R-squared	0.215	0.218	0.219	0.224

Note: Robust pval in parentheses. The dependent variable is annual inflation forecast error (difference between observed and forecasted). The sample refers to euro area countries (19).

 $^{^{\}dagger}$ Real ST R refers to real short term interest rate forecast.

 $^{^{\}ddagger}$ Fiscal refers to fiscal stance forecast.

 $[\]S$ Energy refers to energy price forecast errors.

^{***}p < 0.01, **p < 0.05, *p < 0.1.

BANCO DE ESPAÑA PUBLICATIONS

WORKING PAPERS

- 2401 LAURA HOSPIDO, NAGORE IRIBERRI and MARGARITA MACHELETT: Gender gaps in financial literacy: a multi-arm RCT to break the response bias in surveys.
- 2402 RUBÉN DOMÍNGUEZ-DÍAZ, SAMUEL HURTADO and CAROLINA MENÉNDEZ: The medium-term effects of investment stimulus
- 2403 CLODOMIRO FERREIRA, JOSÉ MIGUEL LEIVA, GALO NUÑO, ÁLVARO ORTIZ, TOMASA RODRIGO and SIRENIA VAZQUEZ: The heterogeneous impact of inflation on households' balance sheets.
- 2404 JORGE ABAD, GALO NUÑO and CARLOS THOMAS: CBDC and the operational framework of monetary policy.
- 2405 STÉPHANE BONHOMME and ANGELA DENIS: Estimating individual responses when tomorrow matters.
- 2406 LAURA ÁLVAREZ-ROMÁN, SERGIO MAYORDOMO, CARLES VERGARA-ALERT and XAVIER VIVES: Climate risk, soft information and credit supply.
- 2407 JESÚS FERNÁNDEZ-VILLAVERDE, JOËL MARBET, GALO NUÑO and OMAR RACHEDI: Inequality and the zero lower
- 2408 PABLO BURRIEL, MAR DELGADO-TÉLLEZ, CAMILA FIGUEROA, IVÁN KATARYNIUK and JAVIER J. PÉREZ: Estimating the contribution of macroeconomic factors to sovereign bond spreads in the euro area.
- 2409 LUIS E. ROJAS and DOMINIK THALER: The bright side of the doom loop: banks' sovereign exposure and default incentives.
- 2410 SALOMÓN GARCÍA-VILLEGAS and ENRIC MARTORELL: Climate transition risk and the role of bank capital requirements
- 2411 MIKEL BEDAYO and JORGE E. GALÁN: The impact of the Countercyclical Capital Buffer on credit: Evidence from its accumulation and release before and during COVID-19.
- 2412 EFFROSYNI ADAMOPOULOU, LUIS DÍEZ-CATALÁN and ERNESTO VILLANUEVA: Staggered contracts and unemployment during recessions.
- 2413 LUIS FÉRNANDEZ LAFUERZA and JORGE E. GALÁN: Should macroprudential policy target corporate lending? Evidence from credit standards and defaults.
- 2414 STÉPHANE BONHOMME and ANGELA DENIS: Estimating heterogeneous effects: applications to labor economics.
- 2415 LUIS GUIROLA, LAURA HOSPIDO and ANDREA WEBER: Family and career: An analysis across Europe and North America.
- 2416 GERALD P. DWYER, BILJANA GILEVSKA, MARÍA J. NIETO and MARGARITA SAMARTÍN: The effects of the ECB's unconventional monetary policies from 2011 to 2018 on banking assets.
- 2417 NICOLÁS FORTEZA, ELVIRA PRADES and MARC ROCA: Analysing the VAT cut pass-through in Spain using webscraped supermarket data and machine learning.
- 2418 JOSÉ-ELÍAS GALLEGOS: HANK beyond FIRE: Amplification, forward guidance, and belief shocks.
- 2419 DANIEL ALONSO: Stabilisation properties of a SURE-like European unemployment insurance.
- 2420 FRANCISCO GONZÁLEZ, JOSÉ E. GUTIÉRREZ and JOSÉ MARÍA SERENA: Shadow seniority? Lending relationships and borrowers' selective default.
- 2421 ROBERTO BLANCO, MIGUEL GARCÍA-POSADA, SERGIO MAYORDOMO and MARÍA RODRÍGUEZ-MORENO: Access to credit and firm survival during a crisis: the case of zero-bank-debt firms.
- 2422 FERNANDO CEREZO, PABLO GIRÓN, MARÍA T. GONZÁLEZ-PÉREZ and ROBERTO PASCUAL: The impact of sovereign debt purchase programmes. A case study: the Spanish-to-Portuguese bond yield spread.
- 2423 EDGAR SILGADO-GÓMEZ: Sovereign uncertainty.
- 2424 CLODOMIRO FERREIRA, JULIO GÁLVEZ and MYROSLAV PIDKUYKO: Housing tenure, consumption and household debt: life-cycle dynamics during a housing bust in Spain.
- 2425 RUBÉN DOMÍNGUEZ-DÍAZ and SAMUEL HURTADO: Green energy transition and vulnerability to external shocks.
- 2426 JOSEP GISBERT and JOSÉ E. GUTIÉRREZ: Bridging the gap? Fintech and financial inclusion.
- 2427 RODOLFO G. CAMPOS, MARIO LARCH, JACOPO TIMINI, ELENA VIDAL and YOTO V. YOTOV: Does the WTO Promote Trade? A Meta-analysis.
- 2428 SONER BASKAYA, JOSÉ E. GUTIÉRREZ, JOSÉ MARÍA SERENA and SERAFEIM TSOUKAS: Bank supervision and non-performing loan cleansing.
- 2429 TODD E. CLARK, GERGELY GANICS, and ELMAR MERTENS: Constructing fan charts from the ragged edge of SPF forecasts
- 2430 MIGUEL GARCÍA-POSADA and PETER PAZ: The transmission of monetary policy to credit supply in the euro area.
- 2431 KLODIANA ISTREFI, FLORENS ODENDAHL and GIULIA SESTIERI: ECB communication and its impact on financial markets.

- 2432 FRUCTUOSO BORRALLO, LUCÍA CUADRO-SÁEZ, CORINNA GHIRELLI and JAVIER J. PÉREZ: "El Niño" and "La Niña": Revisiting the impact on food commodity prices and euro area consumer prices.
- 2433 VÍCTOR CABALLERO, CORINNA GHIRELLI, ÁNGEL LUIS GÓMEZ and JAVIER J. PÉREZ: The public-private wage GAP in the euro area a decade after the sovereign debt crisis.
- 2434 LIDIA CRUCES, ISABEL MICÓ-MILLÁN and SUSANA PÁRRAGA: Female financial portfolio choices and marital property regimes.
- 2435 RODOLFO G. CAMPOS, ANA-SIMONA MANU, LUIS MOLINA and MARTA SUÁREZ-VARELA: China's financial spillovers to emerging markets.
- 2436 LUDOVIC PANON, LAURA LEBASTARD, MICHELE MANCINI, ALESSANDRO BORIN, PEONARE CAKA, GIANMARCO CARIOLA, DENNIS ESSERS, ELENA GENTILI, ANDREA LINARELLO, TULLIA PADELLINI, FRANCISCO REQUENA and JACOPO TIMINI: Inputs in Distress: Geoeconomic Fragmentation and Firms' Sourcing.
- 2437 DANIEL DEJUAN-BITRIA, WAYNE R. LANDSMAN, SERGIO MAYORDOMO and IRENE ROIBÁS: How do changes in financial reporting standards affect relationship lending?
- 2438 ALICIA AGUILAR and RICARDO GIMENO: Discrete Probability Forecasts: What to expect when you are expecting a monetary policy decision.
- 2439 RODOLFO G. CAMPOS, JESÚS FERNÁNDEZ-VILLAVERDE, GALO NUÑO and PETER PAZ: Navigating by Falling Stars: Monetary Policy with Fiscally Driven Natural Rates.
- 2440 ALEJANDRO CASADO and DAVID MARTÍNEZ-MIERA: Local lending specialization and monetary policy.
- 2441 JORGE ABAD, DAVID MARTÍNEZ-MIERA and JAVIER SUÁREZ: A macroeconomic model of banks' systemic risk taking.
- 2442 JOSEP PIJOAN-MAS and PAU ROLDAN-BLANCO: Dual labor markets and the equilibrium distribution of firms.
- 2443 OLYMPIA BOVER, LAURA HOSPIDO and ANA LAMO: Gender and Career Progression: Evidence from the Banco de España.
- 2444 JESÚS FERNÁNDEZ-VILLAVERDE, GALO NUÑO and JESSE PERLA: Taming the curse of dimensionality: quantitative economics with deep learning.
- 2445 CLODOMIRO FERREIRA and STEFANO PICA: Households' subjective expectations: disagreement, common drivers and reaction to monetary policy.
- 2446 ISABEL MICÓ-MILLÁN: Inheritance Tax Avoidance Through the Family Firm.
- 2447 MIKEL BEDAYO, EVA VALDEOLIVAS and CARLOS PÉREZ: The stabilizing role of local claims in local currency on the variation of foreign claims.
- 2501 HENRIQUE S. BASSO, MYROSLAV PIDKUYKO and OMAR RACHEDi: Opening the black box: aggregate implications of public investment heterogeneity.
- 2502 MARCO BARDOSCIA, ADRIAN CARRO, MARC HINTERSCHWEIGER, MAURO NAPOLETANO, LILIT POPOYAN, ANDREA ROVENTINI and ARZU ULUC: The impact of prudential regulations on the UK housing market and economy: insights from an agent-based model.
- 2503 IRINA BALTEANU, KATJA SCHMIDT and FRANCESCA VIANI: Sourcing all the eggs from one basket: trade dependencies and import prices.
- 2504 RUBÉN VEIGA DUARTE, SAMUEL HURTADO, PABLO A. AGUILAR GARCÍA, JAVIER QUINTANA GONZÁLEZ and CAROLINA MENÉNDEZ ÁLVAREZ: CATALIST: A new, bigger, better model for evaluating climate change transition risks at Banco de España.
- 2505 PILAR GARCÍA and DIEGO TORRES: Perceiving central bank communications through press coverage.
- 2506 MAR DELGADO-TÉLLEZ, JAVIER QUINTANA and DANIEL SANTABÁRBARA: Carbon pricing, border adjustment and renewable energy investment: a network approach.
- 2507 MARTA GARCÍA RODRÍGUEZ: The role of wage expectations in the labor market.
- 2508 REBECA ANGUREN, GABRIEL JIMÉNEZ and JOSÉ-LUIS PEYDRÓ: Bank capital requirements and risk-taking: evidence from Basel III.
- 2509 JORGE E. GALÁN: Macroprudential policy and the tail risk of credit growth.
- 2510 PETER KARADI, ANTON NAKOV, GALO NUÑO, ERNESTO PASTÉN and DOMINIK THALER: Strike while the Iron is Hot: Optimal Monetary Policy with a Nonlinear Phillips Curve.
- 2511 MATTEO MOGLIANI and FLORENS ODENDAHL: Density forecast transformations.
- 2512 LUCÍA LÓPEZ, FLORENS ODENDAHL, SUSANA PÁRRAGA and EDGAR SILGADO-GÓMEZ: The pass-through to inflation of gas price shocks.
- 2513 CARMEN BROTO and OLIVIER HUBERT: Desertification in Spain: Is there any impact on credit to firms?
- 2514 ANDRÉS ALONSO-ROBISCO, JOSÉ MANUEL CARBÓ, PEDRO JESÚS CUADROS-SOLAS and JARA QUINTANERO: The effects of open banking on fintech providers: evidence using microdata from Spain.
- $2515 \quad \hbox{RODOLFO G. CAMPOS and JACOPO TIMINI: Trade bloc enlargement when many countries join at once.}$
- 2516 CORINNA GHIRELLI, JAVIER J. PÉREZ and DANIEL SANTABÁRBARA: Inflation and growth forecast errors and the sacrifice ratio of monetary policy in the euro area.